

Moderate deviations and laws of the iterated logarithm for the renormalized self-intersection local times of planar random walks

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Abstract

We study moderate deviations for the renormalized self-intersection local time of planar random walks. We also prove laws of the iterated logarithm for such local times

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1 Introduction

Let $\{S_n\}$ be a symmetric random walk on \mathbb{Z}^2 with covariance matrix Γ . Let

$$B_n = \sum_{1 \leq j < k \leq n} \delta(S_j, S_k) \tag{1.1}$$

where

$$\delta(x, y) = \begin{cases} 1 & \text{if } x = y \\ 0 & \text{otherwise} \end{cases} \tag{1.2}$$

is the usual Kroenecker delta. We refer to B_n as the self-intersection local time up to time n . We call

$$\gamma_n =: B_n - \mathbb{E} B_n$$

the renormalized self-intersection local time of the random walk up to time n .

In (13) and (16) it was shown that γ_n , appropriately scaled, converges to the renormalized self-intersection local time of planar Brownian motion. (For a recent almost sure invariance principle see (5).) Renormalized self-intersection local time for Brownian motion was originally studied by Varadhan (18) for its role in quantum field theory. Renormalized self-intersection local time turns out to be the right tool for the solution of certain “classical” problems such as the asymptotic expansion of the area of the Wiener sausage in the plane and the range of random walks, (4), (14), (13).

One of the applications of self-intersection local time is to polymer growth. If S_n is a planar random walk and \mathbb{P} is its law, one can construct self-repelling and self-attracting random walks by defining

$$d\mathbb{Q}_n/d\mathbb{P} = c_n e^{\zeta B_n/n},$$

where ζ is a parameter and c_n is chosen to make \mathbb{Q}_n a probability measure. When $\zeta < 0$, more weight is given to those paths with a small number of self-intersections, hence \mathbb{Q}_n is a model for a self-repelling random walk. When $\zeta > 0$, more weight is given to paths with a large number of self-intersections, leading to a self-attracting random walk. Since $\mathbb{E} B_n$ is deterministic, by modifying c_n , we can write

$$d\mathbb{Q}_n/d\mathbb{P} = c_n e^{\zeta(B_n - \mathbb{E} B_n)/n}.$$

It is known that for small positive ζ the self-attracting random walk grows with n while for large ζ it “collapses”; in the case of collapse its diameter remains bounded in mean square, while in the case of non-collapse the diameter is of order n in mean square. It has been an open problem to determine ζ_c , the critical value of ζ at which the phase transition takes place. The work (2) suggested that the critical value ζ_c could be expressed in terms of the best constant $\kappa(2, 2)$ of a certain Gagliardo-Nirenberg inequality, but that work was for planar Brownian motion, not for random walks. In (2) it was shown that $\mathbb{E} e^{\zeta \tilde{\gamma}_1}$ is finite or infinite according to whether ζ is less than or greater than $\kappa(2, 2)^{-4}$, where $\tilde{\gamma}_1$ is the renormalized self-intersection time for planar Brownian motion. In the current paper we obtain moderate deviations estimates for γ_n and these are in terms of the best constant of the Gagliardo-Nirenberg inequality; see Theorem 1.1. However the critical constant ζ_c is different from $\kappa(2, 2)^{-4}$ (see Remark 1.4) and it is still an open problem to determine it. See (6) and (7) for details and further information on these models.

In the present paper we study moderate deviations of γ_n . Before stating our main theorem we recall one of the Gagliardo-Nirenberg inequalities:

$$\|f\|_4 \leq C \|\nabla f\|_2^{1/2} \|f\|_2^{1/2}, \quad (1.3)$$

which is valid for $f \in C^1$ with compact support, and can then be extended to

$$W^{1,2}(R^2) =: \{f \in L^2(R^2) \mid \nabla f \in L^2(R^2)\} \quad (1.4)$$

We define $\kappa(2, 2)$ to be the best constant in (1.3), that is,

$$\kappa(2, 2) =: \inf \left\{ C > 0 \mid \|f\|_4 \leq C \|\nabla f\|_2^{1/2} \|f\|_2^{1/2}, \quad \forall f \in W^{1,2}(R^2) \right\} \quad (1.5)$$

In particular, $0 < \kappa(2, 2) < \infty$. We note for later reference that

$$\sup_{g \in \mathcal{F}_2} \left\{ \left(\int_{\mathbb{R}^2} |g(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \nabla g \rangle dx \right\} = \frac{1}{2} \kappa^4(2, 2) \quad (1.6)$$

where

$$\mathcal{F}_2 = \{g \in W^{1,2}(\mathbb{R}^2) \mid \|g\|_2 = 1\}. \quad (1.7)$$

The identity (1.6) is a special case of (8, Lemma 8.2).

In this paper we will always assume that the smallest group which supports $\{S_n\}$ is \mathbb{Z}^2 . For simplicity we assume further that our random walk is strongly aperiodic. This is needed to get suitable estimates for the transition probability estimates in the proof of Lemma 2.1 and is also used in an essential way in the proof of Theorem 4.1.

Theorem 1.1. *Let $\{b_n\}$ be a positive sequence satisfying*

$$\lim_{n \rightarrow \infty} b_n = \infty \quad \text{and} \quad b_n = o(n). \quad (1.8)$$

For any $\lambda > 0$,

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{P} \left\{ B_n - \mathbb{E} B_n \geq \lambda n b_n \right\} = -\lambda \sqrt{\det \Gamma} \kappa(2, 2)^{-4}. \quad (1.9)$$

We call Theorem 1.1 a moderate deviations theorem rather than a large deviations result because of the second restriction in (1.8). Our techniques do not apply when this restriction is not present, and in fact it is not hard to show that the value on the right hand side of (1.9) should be different when $b_n \approx n$; see Remark 1.4.

Moderate deviations for $-\gamma_n$ are more subtle. In the next theorem we obtain the correct rate, but not the precise constant.

Theorem 1.2. *Suppose $\mathbb{E} |S_1|^{2+\delta} < \infty$ for some $\delta > 0$. There exist $C_1, C_2 > 0$ such that for any sequence $b_n \rightarrow \infty$ with $b_n = o(n)$*

$$\begin{aligned} -C_1 &\leq \liminf_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\ &\leq \limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\ &\leq -C_2. \end{aligned} \quad (1.10)$$

Here are the corresponding laws of the iterated logarithm for γ_n .

Theorem 1.3.

$$\limsup_{n \rightarrow \infty} \frac{B_n - \mathbb{E} B_n}{n \log \log n} = \det(\Gamma)^{-1/2} \kappa(2, 2)^4 \quad a.s. \quad (1.11)$$

and if $\mathbb{E} |S_1|^{2+\delta} < \infty$ for some $\delta > 0$,

$$\liminf_{n \rightarrow \infty} \frac{B_n - \mathbb{E} B_n}{n \log \log \log n} = -(2\pi)^{-1} \det(\Gamma)^{-1/2} \quad a.s. \quad (1.12)$$

In this paper we deal exclusively with the case where the dimension d is 2. We note that in dimension 1 no renormalization is needed, which makes the results much simpler. See (15; 9). When $d \geq 3$, the renormalized intersection local time is in the domain of attraction of a centered normal random variable. Consequently the tails of the weak limit are expected to be of Gaussian type, and in particular, the tails are symmetric; see (13). As far as we know, the interesting question of moderate deviations in dimensions 3 and larger is still open.

Theorems 1.1-1.3 are the analogues of the theorems proved in (2) for the renormalized self-intersection local time of planar Brownian motion. Although the proofs for the random walk case have some elements in common with those for Brownian motion, the random walk case is considerably more difficult. The major difficulty is the fact that we do not have Gaussian random variables. Consequently, the argument for the lower bound of Theorem 1.1 needs to be very different from the one given in (2, Lemma 3.4). This requires several new tools, such as Theorem 4.1, which we expect will have applications beyond the specific needs of this paper.

Remark 1.4. Without the restriction that $b_n = o(n)$, Theorem 1.1 is not true. To see this, let N be an arbitrarily large integer, let $\varepsilon = 2/N^2$, and let X_i be an i.i.d. sequence of random vectors in \mathbb{Z}^2 that take the values $(N, 0)$, $(-N, 0)$, $(0, N)$, and $(0, -N)$ with probability $\varepsilon/4$ and $\mathbb{P}(X_1 = (0, 0)) = 1 - \varepsilon$. The covariance matrix of the X_i will be the identity. Let $b_n = (1 - \varepsilon)n$. Then the event that $S_i = S_0$ for all $i \leq n$ will have probability at least $(1 - \varepsilon)^n$, and on this event $B_n = n(n - 1)/2$. This shows that

$$\log \mathbb{P}(B_n - \mathbb{E} B_n > nb_n/2) \geq n \log(1 - \varepsilon),$$

which would contradict (1.4).

The same example shows that the critical constant ζ_c in the polymer model is different than the one in (2). We have

$$\mathbb{E} \exp \left\{ C \frac{B_n - \mathbb{E} B_n}{n} \right\} \geq \exp \left\{ -C \frac{\mathbb{E} B_n}{n} \right\} (1 - \varepsilon)^n \exp \left\{ C \frac{n - 1}{2} \right\}.$$

This shows that ζ_c is no more than $2 \log \frac{1}{1 - \varepsilon}$. On the other hand, if ε is sufficiently small, $2 \log \frac{1}{1 - \varepsilon} < \kappa(2, 2)^{-4}$.

This paper is organized as follows. In Section 2 we establish some estimates which are used throughout the paper. Section 3 begins the proof of Theorem 1.1, while a crucial element of that proof, Theorem 4.1, is established in Section 4. Sections 5 and 6 prove the upper and lower bounds for Theorem 1.2, and Section 7 is devoted to the laws of the iterated logarithm.

2 Preliminary Estimates

Let $\{S'_n\}$ be an independent copy of the random walk $\{S_n\}$. Let

$$I_{m,n} = \sum_{j=1}^m \sum_{k=1}^n \delta(S_j, S'_k) \quad (2.1)$$

and set $I_n = I_{n,n}$. Thus

$$I_n = \#\{(j, k) \in [1, n]^2; S_j = S'_k\}. \quad (2.2)$$

Lemma 2.1.

$$\mathbb{E} I_{m,n} \leq c((m+n) \log(m+n) - m \log m - n \log n). \quad (2.3)$$

In particular

$$\mathbb{E}(I_n) \leq cn. \quad (2.4)$$

We also have

$$\mathbb{E} I_{m,n} \leq c\sqrt{mn}. \quad (2.5)$$

Proof Using symmetry and independence

$$\begin{aligned} \mathbb{E} I_{m,n} &= \sum_{j=1}^m \sum_{k=1}^n \mathbb{E} \delta(S_j, S'_k) \\ &= \sum_{j=1}^m \sum_{k=1}^n \mathbb{E} \delta(S_j - S'_k, 0) \\ &= \sum_{j=1}^m \sum_{k=1}^n \mathbb{E} \delta(S_{j+k}, 0) = \sum_{j=1}^m \sum_{k=1}^n p_{j+k}(0) \end{aligned} \quad (2.6)$$

where $p_n(a) = P(S_n = a)$. By (17, p. 75),

$$p_m(0) = \frac{1}{2\pi\sqrt{\det \Gamma}} \frac{1}{m} + o\left(\frac{1}{m}\right) \quad (2.7)$$

so that

$$\mathbb{E} I_{m,n} \leq c \sum_{j=1}^m \sum_{k=1}^n \frac{1}{j+k} \leq c \int_{r=0}^m \int_{s=0}^n \frac{1}{r+s} dr ds \quad (2.8)$$

and (2.3) follows. (2.4) is then immediate. (2.5) follows from (2.8) and the bound $(r+s)^{-1} \leq (\sqrt{rs})^{-1}$. \square

It follows from the proof of (8, Lemma 5.2) that for any integer $k \geq 1$

$$\mathbb{E}(I_n^k) \leq (k!)^2 2^k (1 + \mathbb{E}(I_n))^k. \quad (2.9)$$

Furthermore, by (13, (5.k)) we have that I_n/n converges in distribution to a random variable with finite moments. Hence for any integer $k \geq 1$

$$\lim_{n \rightarrow \infty} \frac{\mathbb{E}(I_n^k)}{n^k} = c_k < \infty. \quad (2.10)$$

Lemma 2.2. *There is a constant $c > 0$ such that*

$$\sup_n \mathbb{E} \exp \left\{ \frac{c}{n} I_n \right\} < \infty. \quad (2.11)$$

Proof. We need only to show that there is a $C > 0$ such that

$$\mathbb{E} I_n^m \leq C^m m! n^m \quad m, n \geq 1.$$

We first consider the case $m \leq n$ and write $l(m, n) = \lfloor n/m \rfloor + 1$. Using (8, Theorem 5.1) with $p = 2$ and $a = m$, and then (2.4), (2.9) and (2.10), we obtain

$$\begin{aligned} (\mathbb{E} I_n^m)^{1/2} &\leq \sum_{\substack{k_1 + \dots + k_m = m \\ k_1, \dots, k_m \geq 0}} \frac{m!}{k_1! \dots k_m!} (\mathbb{E} I_{l(m,n)}^{k_1})^{1/2} \dots (\mathbb{E} I_{l(m,n)}^{k_m})^{1/2} \\ &\leq \sum_{\substack{k_1 + \dots + k_m = m \\ k_1, \dots, k_m \geq 0}} \frac{C^m m!}{k_1! \dots k_m!} k_1! \dots k_m! (\mathbb{E} I_{l(m,n)})^{k_1/2} \dots (\mathbb{E} I_{l(m,n)})^{k_m/2} \\ &\leq \binom{2m-1}{m} m! C^m \left(\frac{n}{m}\right)^{m/2} \leq \binom{2m}{m} m! C^m \left(\frac{n}{m}\right)^{m/2} \end{aligned} \quad (2.12)$$

where $C > 0$ can be chosen independently of m and n . Hence

$$\mathbb{E} I_n^m \leq \binom{2m}{m}^2 C^m (m!)^2 \left(\frac{n}{m}\right)^m \leq \binom{2m}{m}^2 C^m m! n^m. \quad (2.13)$$

Notice that

$$\binom{2m}{m} \leq 4^m. \quad (2.14)$$

For the case $m > n$, notice that $I_n \leq n^2$. Trivially,

$$\mathbb{E} I_n^m \leq n^{2m} \leq m^m n^m \leq C^m m! n^m,$$

where the last step follows from Stirling's formula. \square

For any random variable X we define

$$\overline{X} =: X - \mathbb{E} X.$$

We write

$$(m, n]_{<}^2 = \{(j, k) \in (m, n]_{<}^2; \quad j < k\} \quad (2.15)$$

For any $A \subset \{(j, k) \in (\mathbb{Z}^+)^2; \quad j < k\}$, write

$$B(A) = \sum_{(j,k) \in A} \delta(S_j, S_k) \quad (2.16)$$

In our proofs we will use several decompositions of B_n . If J_1, \dots, J_ℓ are consecutive disjoint blocks of integers whose union is $\{1, \dots, n\}$, we have

$$B_n = \sum_i B((J_i \times J_i) \cap (0, n]_{<}^2) + \sum_{i < j} B(J_i \times J_j)$$

and also

$$B_n = \sum_i B((J_i \times J_i) \cap (0, n]_{<}^2) + \sum_i B(\cup_{j=1}^{i-1} J_j \times J_i).$$

Lemma 2.3. *There is a constant $c > 0$ such that*

$$\sup_n \mathbb{E} \exp \left\{ \frac{c}{n} |\overline{B}_n| \right\} < \infty. \quad (2.17)$$

Proof. We first prove that there is $c > 0$ such that

$$M \equiv \sup_n \mathbb{E} \exp \left\{ \frac{c}{2^n} |\overline{B}_{2^n}| \right\} < \infty. \quad (2.18)$$

We have

$$\begin{aligned} & B_{2^n} \\ &= \sum_{j=1}^n \sum_{k=1}^{2^{j-1}} B \left(((2k-2)2^{n-j}, (2k-1)2^{n-j}] \times ((2k-1)2^{n-j}, (2k)2^{n-j}] \right). \end{aligned} \quad (2.19)$$

Write

$$\begin{aligned} \alpha_{j,k} &= B \left(((2k-2)2^{n-j}, (2k-1)2^{n-j}] \times ((2k-1)2^{n-j}, (2k)2^{n-j}] \right) \\ &\quad - \mathbb{E} B \left(((2k-2)2^{n-j}, (2k-1)2^{n-j}] \times ((2k-1)2^{n-j}, (2k)2^{n-j}] \right). \end{aligned} \quad (2.20)$$

For each $1 \leq j \leq n$, the random variables $\alpha_{j,k}$, $k = 1, \dots, 2^{j-1}$ are i.i.d. with common distribution $I_{2^{n-j}} - \mathbb{E} I_{2^{n-j}}$. By the previous lemma there exists $\delta > 0$ such that

$$\sup_n \sup_{j \leq n} \mathbb{E} \exp \left\{ \delta \frac{1}{2^{n-j}} |\alpha_{j,1}| \right\} < \infty. \quad (2.21)$$

By (3, Lemma 1), there exists $\theta > 0$ such that

$$\begin{aligned} C(\theta) &\equiv \sup_n \sup_{j \leq n} \mathbb{E} \exp \left\{ \theta 2^{j/2} \frac{1}{2^n} \left| \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} \\ &= \sup_n \sup_{j \leq n} \mathbb{E} \exp \left\{ \theta 2^{-j/2} \frac{1}{2^{n-j}} \left| \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} < \infty. \end{aligned} \quad (2.22)$$

Write

$$\lambda_N = \prod_{j=1}^N (1 - 2^{-j/2}) \quad \text{and} \quad \lambda_\infty = \prod_{j=1}^{\infty} (1 - 2^{-j/2}). \quad (2.23)$$

Using Hölder's inequality with $1/p = 1 - 2^{-n/2}$, $1/q = 2^{-n/2}$ we have

$$\begin{aligned}
& \mathbb{E} \exp \left\{ \lambda_n \frac{\theta}{2^n} \left| \sum_{j=1}^n \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} \\
& \leq \left(\mathbb{E} \exp \left\{ \lambda_{n-1} \frac{\theta}{2^n} \left| \sum_{j=1}^{n-1} \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} \right)^{1-2^{-n/2}} \\
& \quad \times \left(\mathbb{E} \exp \left\{ 2^{n/2} \lambda_n \frac{\theta}{2^n} \left| \sum_{k=1}^{2^{n-1}} \alpha_{n,k} \right| \right\} \right)^{2^{-n/2}} \\
& \leq \mathbb{E} \exp \left\{ \lambda_{n-1} \frac{\theta}{2^n} \left| \sum_{j=1}^{n-1} \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} C(\theta)^{2^{-n/2}}.
\end{aligned} \tag{2.24}$$

Repeating this procedure,

$$\begin{aligned}
& \mathbb{E} \exp \left\{ \lambda_n \frac{\theta}{2^n} \left| \sum_{j=1}^n \sum_{k=1}^{2^{j-1}} \alpha_{j,k} \right| \right\} \\
& \leq C(\theta)^{2^{-1/2} + \dots + 2^{-n/2}} \leq C(\theta)^{2^{-1/2}(1-2^{-1/2})^{-1}}.
\end{aligned} \tag{2.25}$$

So we have

$$\sup_n \mathbb{E} \exp \left\{ \lambda_\infty \frac{\theta}{2^n} \left| \bar{B}_{2^n} \right| \right\} < \infty. \tag{2.26}$$

We now prove our lemma for general n . Given an integer $n \geq 2$, we have the following unique representation:

$$n = 2^{m_1} + 2^{m_2} + \dots + 2^{m_l} \tag{2.27}$$

where $m_1 > m_2 > \dots > m_l \geq 0$ are integers. Write

$$n_0 = 0 \text{ and } n_i = 2^{m_1} + \dots + 2^{m_i}, \quad i = 1, \dots, l. \tag{2.28}$$

Then

$$\begin{aligned}
\sum_{1 \leq j < k \leq n} \delta(S_j, S_k) &= \sum_{i=1}^l \sum_{n_{i-1} < j < k \leq n_i} \delta(S_j, S_k) + \sum_{i=1}^{l-1} B((n_{i-1}, n_i] \times (n_i, n]) \\
&=: \sum_{i=1}^l B_{2^{m_i}}^{(i)} + \sum_{i=1}^{l-1} A_i.
\end{aligned} \tag{2.29}$$

By Hölder's inequality, with M as in (2.18)

$$\begin{aligned}
& \mathbb{E} \exp \left\{ \frac{c}{n} \left| \sum_{i=1}^l (B_{2^{m_i}}^{(i)} - \mathbb{E} B_{2^{m_i}}^{(i)}) \right| \right\} \\
& \leq \prod_{i=1}^l \left(\mathbb{E} \exp \left\{ \frac{c}{2^{m_i}} \left| B_{2^{m_i}}^{(i)} - \mathbb{E} B_{2^{m_i}}^{(i)} \right| \right\} \right)^{\frac{2^{m_i}}{n}} \leq \prod_{i=1}^l M^{2^{m_i}/n} = M.
\end{aligned} \tag{2.30}$$

Using Hölder's inequality,

$$\mathbb{E} \exp \left\{ \frac{c}{n} \sum_{i=1}^{l-1} A_i \right\} \leq \prod_{i=1}^{l-1} \left(\mathbb{E} \exp \left\{ \frac{c}{2^{m_i}} A_i \right\} \right)^{\frac{2^{m_i}}{n}}. \quad (2.31)$$

Notice that for each $1 \leq i \leq l-1$,

$$A_i \stackrel{d}{=} \sum_{j=1}^{2^{m_i}} \sum_{k=1}^{n-n_i} \delta(S_j, S'_k) \leq \sum_{j=1}^{2^{m_i}} \sum_{k=1}^{2^{m_i}} \delta(S_j, S'_k), \quad (2.32)$$

where the inequality follows from

$$n - n_i = 2^{m_{i+1}} + \dots + 2^{m_l} \leq 2^{m_i}. \quad (2.33)$$

Using (2.32) and Lemma 2.1, we can take $c > 0$ so that

$$\mathbb{E} \exp \left\{ \frac{c}{2^{m_i}} A_i \right\} \leq \sup_n \mathbb{E} \exp \left\{ \frac{c}{n} I_n \right\} \equiv N < \infty. \quad (2.34)$$

Consequently,

$$\mathbb{E} \exp \left\{ \frac{c}{n} \sum_{i=1}^{l-1} A_i \right\} \leq \prod_{i=1}^{l-1} N^{2^{m_i}/n} \leq N. \quad (2.35)$$

In particular, this shows that

$$\mathbb{E} \left\{ \frac{c}{n} \sum_{i=1}^{l-1} A_i \right\} \leq N. \quad (2.36)$$

Combining (2.35) and (2.36) with (2.30) we have

$$\sup_n \mathbb{E} \exp \left\{ \frac{c}{2n} |\bar{B}_n| \right\} < \infty. \quad (2.37)$$

□

Lemma 2.4.

$$\mathbb{E} B_n = \frac{1}{2\pi\sqrt{\det \Gamma}} n \log n + o(n \log n), \quad (2.38)$$

and if $\mathbb{E} |S_1|^{2+2\delta} < \infty$ for some $\delta > 0$ then

$$\mathbb{E} B_n = \frac{1}{2\pi\sqrt{\det \Gamma}} n \log n + O(n). \quad (2.39)$$

Proof.

$$\mathbb{E} B_n = \mathbb{E} \sum_{1 \leq j < k \leq n} \delta(S_j, S_k) = \sum_{1 \leq j < k \leq n} p_{k-j}(0) \quad (2.40)$$

where $p_m(x) = \mathbb{E}(S_m = x)$. If $\mathbb{E} |S_1|^{2+2\delta} < \infty$, then by (12, Proposition 6.7),

$$p_m(0) = \frac{1}{2\pi\sqrt{\det \Gamma}} \frac{1}{m} + o\left(\frac{1}{m^{1+\delta}}\right). \quad (2.41)$$

Since the last term is summable, it will contribute $O(n)$ to (2.40). Also,

$$\sum_{1 \leq j < k \leq n} \frac{1}{k-j} = \sum_{m=1}^n \sum_{i=1}^{n-m} \frac{1}{m} = \sum_{m=1}^n \frac{n-m}{m} = n \sum_{m=1}^n \frac{1}{m} - n \quad (2.42)$$

and our Lemma follows from the well known fact that

$$\sum_{m=1}^n \frac{1}{m} = \log n + \gamma + O\left(\frac{1}{n}\right) \quad (2.43)$$

where γ is Euler's constant.

If we only assume finite second moments, instead of (2.41) we use (2.7) and proceed as above. \square

Lemma 2.5. *For any $\theta > 0$*

$$\sup_n \mathbb{E} \exp \left\{ \frac{\theta}{n} (\mathbb{E} B_n - B_n) \right\} < \infty \quad (2.44)$$

and for any $\lambda > 0$

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq \lambda n b_n \right\} = -\infty. \quad (2.45)$$

Proof. By Lemma 2.3 this is true for some $\theta_o > 0$. For any $\theta > \theta_o$, take an integer $m \geq 1$ such that $\theta m^{-1} < \theta_o$. We can write any n as $n = rm + i$ with $1 \leq i < m$. Then

$$\begin{aligned} & \mathbb{E} B_n - B_n \\ & \leq \sum_{j=1}^m \left[\mathbb{E} \sum_{(j-1)r < k, l \leq jr} \delta(S_k, S_l) - \sum_{(j-1)r < k, l \leq jr} \delta(S_k, S_l) \right] + \mathbb{E} B_n - m \mathbb{E} B_r. \end{aligned} \quad (2.46)$$

We claim that

$$\mathbb{E} B_n - m \mathbb{E} B_r = O(n). \quad (2.47)$$

To see this, write

$$\mathbb{E} B_n - m \mathbb{E} B_r = \mathbb{E} B_n - \sum_{l=1}^m \mathbb{E} B((l-1)r, lr]_{<}^2) \quad (2.48)$$

Notice that

$$\begin{aligned} & B_n - \sum_{l=1}^m B((l-1)r, lr]_{<}^2) \\ & = \sum_{l=1}^m B(((l-1)r, lr] \times (lr, mr]) + B((mr, n]_{<}^2) \\ & \quad + B((0, mr] \times (mr, n]) \end{aligned} \quad (2.49)$$

Since

$$B(((l-1)r, lr] \times (lr, mr]) \stackrel{d}{=} I_{r, (m-l)r} \quad (2.50)$$

by (2.3) we have

$$\begin{aligned} & \mathbb{E} B((l-1)r, lr] \times (lr, mr]) \\ & \leq C \left\{ (m - (l-1))r \log(m - (l-1))r \right. \\ & \qquad \qquad \qquad \left. - ((m-l)r \log((m-l)r) - r \log r) \right\} \end{aligned} \tag{2.51}$$

Therefore

$$\begin{aligned} & \sum_{l=1}^m \mathbb{E} B((l-1)r, lr] \times (lr, mr]) \\ & \leq C \sum_{l=1}^m \left\{ (m - (l-1))r \log(m - (l-1))r \right. \\ & \qquad \qquad \qquad \left. - ((m-l)r \log((m-l)r) - r \log r) \right\} \\ & = C \left\{ mr \log mr - mr \log r \right\} = Cmr \log m. \end{aligned} \tag{2.52}$$

Using (2.5) for $\mathbb{E} B((0, mr] \times (mr, n]) = \mathbb{E} I_{mr,i}$ and (2.38) for $\mathbb{E} B((mr, n]_{<}^2)$ then completes the proof of (2.47).

Note that the summands in (2.46) are independent. Therefore, for some constant $C > 0$ depending only on θ and m ,

$$\mathbb{E} \exp \left\{ \frac{\theta}{n} (\mathbb{E} B_n - B_n) \right\} \leq C \left(\mathbb{E} \exp \left\{ \frac{\theta}{n} (\mathbb{E} B_r - B_r) \right\} \right)^m \tag{2.53}$$

which proves (2.44), since $\theta/n \leq \theta/mr < \theta_o/r$ and $r \rightarrow \infty$ as $n \rightarrow \infty$.

Then, by Chebyshev's inequality, for any fixed $h > 0$

$$\mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq \lambda n b_n \right\} \leq e^{-h\lambda b_n} \mathbb{E} \exp \left\{ \frac{h}{n} (\mathbb{E} B_n - B_n) \right\} \tag{2.54}$$

so that by (2.44)

$$\limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq \lambda n b_n \right\} \leq -h\lambda. \tag{2.55}$$

Since $h > 0$ is arbitrary, this proves (2.45). □

3 Proof of Theorem 1.1

By the Gärtner-Ellis theorem ((11, Theorem 2.3.6)), we need only prove

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B_n - \mathbb{E} B_n|^{1/2} \right\} = \frac{1}{4} \kappa(2, 2)^4 \theta^2 \det(\Gamma)^{-1/2}. \tag{3.1}$$

Indeed, by the Gärtner-Ellis theorem the above implies that

$$\lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{P} \left\{ |B_n - \mathbb{E} B_n| \geq \lambda n b_n \right\} = -\lambda \sqrt{\det(\Gamma)} \kappa(2, 2)^{-4}. \tag{3.2}$$

Using (2.45) we will then have Theorem 1.1. It thus remains to prove (3.1).

Let f be a symmetric probability density function in the Schwartz space $\mathcal{S}(\mathbb{R}^2)$ of C^∞ rapidly decreasing functions. Let $\epsilon > 0$ and write

$$f_\epsilon(x) = \epsilon^{-2} f(\epsilon^{-1}x), \quad x \in \mathbb{R}^2 \quad (3.3)$$

and

$$l(n, x) = \sum_{k=1}^n \delta(S_k, x), \quad l(n, x, \epsilon) = \sum_{k=1}^n f_{\epsilon(b_n^{-1}n)^{1/2}}(S_k - x). \quad (3.4)$$

$l(n, x)$ is the local time at x , that is, the number of visits to x up till time n . $l(n, x, \epsilon)$ is a smoothed version of the local time. Note that

$$\frac{1}{2} \sum_x l^2(n, x) = \frac{1}{2} \sum_{i,j=1}^n \delta(S_i, S_j) = B_n + \frac{1}{2}n. \quad (3.5)$$

Hence we can replace B_n in (3.1) by $\frac{1}{2} \sum_x l^2(n, x)$. This motivates the next Theorem, proved below, which shows that in certain sense $B_n - \mathbb{E} B_n$ is close to $\frac{1}{2} \sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon)$.

Theorem 3.1. *For any $\theta > 0$,*

$$\lim_{\epsilon \rightarrow 0} \lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B_n - \mathbb{E} B_n - \frac{1}{2} \sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon)|^{1/2} \right\} = 0.$$

This Theorem together with a careful use of Hölder's inequality, the details of which are spelled out in the proof of (10, Theorem 1), shows that

$$\begin{aligned} & \lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B_n - \mathbb{E} B_n|^{1/2} \right\} \\ &= \lim_{\epsilon \rightarrow 0} \lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \frac{\theta}{\sqrt{2}} \sqrt{\frac{b_n}{n}} \left(\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon) \right)^{1/2} \right\}. \end{aligned} \quad (3.6)$$

By a minor modification of (8, Theorem 3.1),

$$\begin{aligned} & \lim_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \frac{\theta}{\sqrt{2}} \sqrt{\frac{b_n}{n}} \left(\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon) \right)^{1/2} \right\} \\ &= \sup_{g \in \mathcal{F}_2} \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g^2 * f_\epsilon(x)|^2 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\}. \end{aligned} \quad (3.7)$$

It is easy to see that

$$\begin{aligned} & \lim_{\epsilon \rightarrow 0} \sup_{g \in \mathcal{F}_2} \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g^2 * f_\epsilon(x)|^2 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\} \\ &= \sup_{g \in \mathcal{F}_2} \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\}. \end{aligned} \quad (3.8)$$

The upper bound is immediate since by Young's inequality $\|g^2 * f_\epsilon\|_2 \leq \|g^2\|_2$, whereas the lower bound follows from the fact that for any $g \in \mathcal{F}_2$ the left hand side is greater than

$$\begin{aligned} & \lim_{\epsilon \rightarrow 0} \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g^2 * f_\epsilon(x)|^2 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\} \\ & = \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\}. \end{aligned} \quad (3.9)$$

Furthermore, by writing $g(x) = \frac{\theta \det(\Gamma)^{-1/2}}{\sqrt{2}} f\left(\frac{\theta \det(\Gamma)^{-1/4}}{\sqrt{2}} \Gamma^{-1/2} x\right)$ we see that

$$\begin{aligned} & \sup_{g \in \mathcal{F}_2} \left\{ \frac{\theta}{\sqrt{2}} \left(\int_{\mathbb{R}^2} |g(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g, \Gamma \nabla g \rangle dx \right\} \\ & = \frac{\theta^2}{2} \det(\Gamma)^{-1/2} \sup_{f \in \mathcal{F}_2} \left\{ \left(\int_{\mathbb{R}^2} |f(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla f, \nabla f \rangle dx \right\} \\ & = \frac{\theta^2}{4} \det(\Gamma)^{-1/2} \kappa(2, 2)^4 \end{aligned} \quad (3.10)$$

where the last step used (1.6). (3.1) then follows from (3.6)-(3.10).

Proof of Theorem 3.1: B_n is a double sum over i, j and the same is true of $\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon)$. The basic idea of our proof is that the contribution to B_n in the scale of interest to us from i, j near the diagonal is almost deterministic and therefore will be controlled by $\mathbb{E} B_n$, while the contribution to $\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon)$ from i, j near the diagonal is itself negligible, due to the smoothed out nature of $l(n, x, \epsilon)$. This is the content of Lemma 3.2. The heart of the proof of our Theorem is to show that the contributions to B_n and $\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon)$ from i, j far from the diagonal are 'almost' the same. This is the content of Lemma 3.3 whose proof extends through the following section. In order to get started we need some terminology to formalize the idea of 'near the diagonal' and 'far from the diagonal'.

Let $l > 1$ be a large but fixed integer. Divide $[1, n]$ into l disjoint subintervals D_1, \dots, D_l , each of length $\lfloor n/l \rfloor$ or $\lfloor n/l \rfloor + 1$. Write

$$D_i^* = \{(j, k) \in D_i^2; j < k\} \quad i = 1, \dots, l \quad (3.11)$$

With the notation of (2.16) we have

$$B_n = \sum_{i=1}^l B(D_i^*) + \sum_{1 \leq j < k \leq l} B(D_j \times D_k) \quad (3.12)$$

Define a_j, b_j so that $D_j = (a_j, b_j]$ ($1 \leq j \leq l$). Notice that

$$\begin{aligned} B(D_j \times D_k) & = \sum_{n_1 \in D_j, n_2 \in D_k} \delta(S_{n_1}, S_{n_2}) \\ & = \sum_{n_1 \in D_j, n_2 \in D_k} \delta((S_{n_1} - S_{b_j}) + S_{b_j}, S_{a_k} + (S_{n_2} - S_{a_k})) \\ & = \sum_{n_1 \in D_j, n_2 \in D_k} \delta((S_{n_1} - S_{b_j}), Z + (S_{n_2} - S_{a_k})) \end{aligned} \quad (3.13)$$

with $Z \stackrel{d}{=} S_{a_k} - S_{b_j}$, so that $Z, S_{n_1} - S_{b_j}, S_{n_2} - S_{a_k}$ are independent. Then as in (2.6)

$$\mathbb{E} B(D_j \times D_k) = \mathbb{E} \sum_{n_1 \in D_j, n_2 \in D_k} p_{b_j - n_1 + n_2 - a_k}(Z). \quad (3.14)$$

Note that since X_1 is symmetric its characteristic function $\phi(\lambda)$ is real so that $\phi^2(\lambda) \geq 0$. Thus for any m

$$\sup_x p_{2m}(x) = \sup_x \frac{1}{2\pi} \int e^{i\lambda \cdot x} \phi^{2m}(\lambda) d\lambda \leq \frac{1}{2\pi} \int \phi^{2m}(\lambda) d\lambda = p_{2m}(0) \quad (3.15)$$

and $\sup_x p_{2m+1}(x) = \sup_x \sum_y p_1(y) p_{2m}(x-y) \leq p_{2m}(0)$. Then using (3.14) we have

$$\mathbb{E} B(D_j \times D_k) \leq 2 \sum_{n_1 \in D_j, n_2 \in D_k} (p_{b_j - n_1 + n_2 - a_k}(0) + p_{b_j - n_1 + n_2 - a_k - 1}(0)). \quad (3.16)$$

As in the proof of (2.4) we then have that

$$\mathbb{E} B(D_j \times D_k) \leq cn/l. \quad (3.17)$$

Hence,

$$\begin{aligned} & B_n - \mathbb{E} B_n \\ &= \sum_{i=1}^l [B(D_i^*) - \mathbb{E} B(D_i^*)] + \sum_{1 \leq j < k \leq l} B(D_j \times D_k) - \mathbb{E} \sum_{1 \leq j < k \leq l} B(D_j \times D_k) \\ &= \sum_{i=1}^l [B(D_i^*) - \mathbb{E} B(D_i^*)] + \sum_{1 \leq j < k \leq l} B(D_j \times D_k) + O(n) \end{aligned} \quad (3.18)$$

where the last line follows from (3.17).

Write

$$\xi_i(n, x, \epsilon) = \sum_{k \in D_i} f_{\epsilon(b_n^{-1}n)^{1/2}}(S_k - x). \quad (3.19)$$

Then

$$\sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon) = \sum_{i=1}^l \sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) + 2 \sum_{1 \leq j < k \leq l} \sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon). \quad (3.20)$$

Therefore, by (3.18)

$$\begin{aligned} & \left| (B_n - \mathbb{E} B_n) - \frac{1}{2} \sum_{x \in \mathbb{Z}^2} l^2(n, x, \epsilon) \right| \\ & \leq \sum_{i=1}^l |B(D_i^*) - \mathbb{E} B(D_i^*)| + \frac{1}{2} \sum_{i=1}^l \sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \\ & \quad + \sum_{1 \leq j < k \leq l} \left| B(D_j \times D_k) - \sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon) \right| + O(n). \end{aligned} \quad (3.21)$$

The proof of Theorem 3.1 is completed in the next two lemmas.

Lemma 3.2. For any $\theta > 0$,

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{i=1}^l \sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \\ \leq l^{-1} \frac{1}{2} \kappa(2, 2)^4 \theta^2 \det(\Gamma)^{-1/2} \end{aligned} \quad (3.22)$$

and

$$\limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{i=1}^l |B(D_i^*) - \mathbb{E} B(D_i^*)| \right)^{1/2} \right\} \leq l^{-1} H \theta^2, \quad (3.23)$$

where

$$H = \left(\sup \left\{ \lambda > 0; \sup_n \mathbb{E} \exp \left\{ \lambda \frac{1}{n} |B_n - \mathbb{E} B_n| \right\} < \infty \right\} \right)^{-1}. \quad (3.24)$$

Proof. Replacing θ by θ/\sqrt{l} , n by n/l , and b_n by $b_n^* = b_{ln}$ (notice that $b_{n/l}^* = b_n$)

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \\ = \limsup_{n \rightarrow \infty} \frac{1}{b_{n/l}^*} \log \mathbb{E} \exp \left\{ \frac{\theta}{\sqrt{l}} \sqrt{\frac{b_{n/l}^*}{n/l}} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \end{aligned} \quad (3.25)$$

Applying Jensen's inequality on the right hand side of (3.7),

$$\begin{aligned} \int_{\mathbb{R}^2} |g^2 * f_\epsilon(x)|^2 &= \int_{\mathbb{R}^2} \left[\int_{\mathbb{R}^2} g^2(x-y) f_\epsilon(y) dy \right]^2 dx \\ &\leq \int \int g^4(x-y) f_\epsilon(y) dy dx = \int f_\epsilon(y) \left[\int g^4(x-y) dx \right] dy \\ &= \left[\int g^4(x) dx \right] \int f_\epsilon(y) dy = \int_{\mathbb{R}^2} g^4(y) dy. \end{aligned}$$

Combining the last two displays with (3.7) we have that

$$\begin{aligned} \limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \\ \leq \sup_{g \in \mathcal{F}_2} \left\{ \frac{\theta}{\sqrt{l}} \left(\int_{\mathbb{R}^2} |g(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} \langle \nabla g(x), \Gamma \nabla g(x) \rangle dx \right\} \\ = l^{-1} \theta^2 \det(\Gamma)^{-1/2} \sup_{h \in \mathcal{F}_2} \left\{ \left(\int_{\mathbb{R}^2} |h(x)|^4 dx \right)^{1/2} - \frac{1}{2} \int_{\mathbb{R}^2} |\nabla h(x)|^2 dx \right\} \\ = \frac{1}{2} l^{-1} \det(\Gamma)^{-1/2} \kappa(2, 2)^4 \theta^2, \end{aligned} \quad (3.26)$$

where the third line follows from the substitution $g(x) = \sqrt{|\det(A)|}f(Ax)$ with a 2×2 matrix A satisfying

$$A^T \Gamma A = \frac{1}{2} l^{-1} \theta^2 \det(\Gamma)^{-1/2} \mathbf{I}_2 \quad (3.27)$$

and the last line of (8, Lemma A.2); here \mathbf{I}_2 is the 2×2 identity matrix.

Given $\delta > 0$, there exist $\bar{a}_1 = (a_{1,1}, \dots, a_{1,l}), \dots, \bar{a}_m = (a_{m,1}, \dots, a_{m,l})$ in \mathbb{R}^l such that $|\bar{a}_1| = \dots = |\bar{a}_m| = 1$ and

$$|z| \leq (1 + \delta) \max\{\bar{a}_1 \cdot z, \dots, \bar{a}_m \cdot z\}, \quad z \in \mathbb{R}^l. \quad (3.28)$$

In particular, with

$$z = \left(\left(\sum_{x \in \mathbb{Z}^2} \xi_1^2(n, x, \epsilon) \right)^{1/2}, \dots, \left(\sum_{x \in \mathbb{Z}^2} \xi_l^2(n, x, \epsilon) \right)^{1/2} \right) \quad (3.29)$$

we have

$$\left(\sum_{i=1}^l \sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \leq (1 + \delta) \max_{1 \leq j \leq m} \sum_{i=1}^l a_{j,i} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2}. \quad (3.30)$$

Hence

$$\begin{aligned} & \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{i=1}^l \sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \\ & \leq \sum_{j=1}^m \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} (1 + \delta) \sum_{i=1}^l a_{j,i} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\} \\ & = \sum_{j=1}^m \prod_{i=1}^l \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} (1 + \delta) a_{j,i} \left(\sum_{x \in \mathbb{Z}^2} \xi_i^2(n, x, \epsilon) \right)^{1/2} \right\}, \end{aligned} \quad (3.31)$$

where the last line follows from independence of $\|\xi_i(n, x, \epsilon)\|_{L^2(\mathbb{Z}^2)}$, $i = 1, \dots, l$. Therefore

$$\begin{aligned} & \limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} \left(\sum_{k=1}^l \sum_{x \in \mathbb{Z}^2} \xi_k^2(n, x, \epsilon) \right)^{1/2} \right\} \\ & \leq \max_{1 \leq j \leq m} \frac{1}{2} l^{-1} \kappa(2, 2)^4 (1 + \delta)^2 \theta^2 \left(\sum_{i=1}^l a_{j,i}^2 \right) \\ & = \frac{1}{2} l^{-1} \det(\Gamma)^{-1/2} \kappa(2, 2)^4 (1 + \delta)^2 \theta^2. \end{aligned} \quad (3.32)$$

Letting $\delta \rightarrow 0^+$ proves (3.22).

By the inequality $ab \leq a^2 + b^2$ we have that

$$\begin{aligned} & \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B_n - \mathbb{E} B_n|^{1/2} \right\} \\ & \leq \exp \{ c^2 \theta^2 b_n \} \mathbb{E} \exp \left\{ c^{-2} \frac{1}{n} |B_n - \mathbb{E} B_n| \right\}, \end{aligned} \quad (3.33)$$

and taking $c^{-2} \uparrow H^{-1}$ we see that for any $\theta > 0$,

$$\limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B_n - \mathbb{E} B_n|^{1/2} \right\} \leq H\theta^2. \quad (3.34)$$

Notice that for any $1 \leq i \leq l$,

$$B(D_i^*) - \mathbb{E} B(D_i^*) \stackrel{d}{=} B_{\#(D_i)} - \mathbb{E} B_{\#(D_i)}. \quad (3.35)$$

We have

$$\mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B(D_i^*) - \mathbb{E} B(D_i^*)|^{1/2} \right\} = \mathbb{E} \exp \left\{ \frac{\theta}{\sqrt{l}} \sqrt{\frac{b_n}{n/l}} |B(D_i^*) - \mathbb{E} B(D_i^*)|^{1/2} \right\}.$$

Replacing θ by θ/\sqrt{l} , n by n/l , and b_n by $b_n^* = b_{ln}$ (notice that $b_{n/l}^* = b_n$) gives

$$\limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B(D_i^*) - \mathbb{E} B(D_i^*)|^{1/2} \right\} \leq l^{-1} H\theta^2. \quad (3.36)$$

Thus (3.23) follows by the same argument we used to prove (3.22). \square

Lemma 3.3. *For any $\theta > 0$ and any $1 \leq j < k \leq l$,*

$$\limsup_{\epsilon \rightarrow 0^+} \limsup_{n \rightarrow \infty} \frac{1}{b_n} \log \mathbb{E} \exp \left\{ \theta \sqrt{\frac{b_n}{n}} |B(D_j \times D_k) - \sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon)|^{1/2} \right\} = 0. \quad (3.37)$$

Proof. We will exploit the fact that for $j < k$ the random walk during the time interval D_k is almost independent of its behavior during the time interval D_j . In the next section we will state and prove Lemma 4.1 which is similar to our Lemma but involves objects defined with respect to two independent random walks. In the remainder of this section we reduce the proof of our Lemma to that of Lemma 4.1.

Fix $1 \leq j < k \leq l$ and estimate

$$B(D_j \times D_k) - \sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon). \quad (3.38)$$

Without loss of generality we may assume that $v =: [n/l] = \#(D_j) = \#(D_k)$. For $y \in \mathbb{Z}^2$ set

$$I_n(y) = \sum_{n_1, n_2=1}^n \delta(S_{n_1}, S'_{n_2} + y). \quad (3.39)$$

Note that $I_n = I_n(0)$. By (3.13) we have that

$$B(D_j \times D_k) \stackrel{d}{=} I_v(Z) \quad (3.40)$$

with Z independent of S, S' .

Similarly, we have

$$\begin{aligned}
& \sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon) \\
&= \sum_{x \in \mathbb{Z}^2} \sum_{n_1 \in D_j, n_2 \in D_k} f_{\epsilon(b_n^{-1}n)^{1/2}}(S_{n_1} - x) f_{\epsilon(b_n^{-1}n)^{1/2}}(S_{n_2} - x) \\
&= \sum_{x \in \mathbb{Z}^2} \sum_{n_1 \in D_j, n_2 \in D_k} f_{\epsilon(b_n^{-1}n)^{1/2}}(x) f_{\epsilon(b_n^{-1}n)^{1/2}}(S_{n_2} - S_{n_1} - x) \\
&= \sum_{n_1 \in D_j, n_2 \in D_k} f_{\epsilon(b_n^{-1}n)^{1/2}} \otimes f_{\epsilon(b_n^{-1}n)^{1/2}}(S_{n_2} - S_{n_1}) \\
&= \sum_{n_1 \in D_j, n_2 \in D_k} f_{\epsilon(b_n^{-1}n)^{1/2}} \otimes f_{\epsilon(b_n^{-1}n)^{1/2}}((S_{n_2} - S_{a_k}) - (S_{n_1} - S_{b_j}) + Z) \quad (3.41)
\end{aligned}$$

where

$$f \otimes f(y) = \sum_{x \in \mathbb{Z}^2} f(x) f(y - x) \quad (3.42)$$

denotes convolution in $L^1(\mathbb{Z}^2)$. It is clear that if $f \in \mathcal{S}(\mathbb{R}^2)$ so is $f \otimes f$. For $y \in \mathbb{Z}^2$, define the link

$$L_{n, \epsilon}(y) = \sum_{n_1, n_2=1}^n f_{\epsilon} \otimes f_{\epsilon}(S'_{n_2} - S_{n_1} + y). \quad (3.43)$$

By (3.41) we have that

$$\sum_{x \in \mathbb{Z}^2} \xi_j(n, x, \epsilon) \xi_k(n, x, \epsilon) \stackrel{d}{=} L_{n, (b_n^{-1}n)^{1/2} \epsilon}(Z) \quad (3.44)$$

with Z independent of S, S' .

Lemma 3.4. *Let $f \in \mathcal{S}(\mathbb{R}^2)$ with Fourier transform \widehat{f} supported on $(-\pi, \pi)^2$. Then for any $r \geq 1$*

$$\int e^{-i\lambda y} (f_r \otimes f_r)(y) dy = (\widehat{f}(r\lambda))^2, \quad \forall \lambda \in \mathbb{R}^2. \quad (3.45)$$

Proof. We have

$$\begin{aligned}
\int e^{-i\lambda y} (f \otimes f)(y) dy &= \sum_{x \in \mathbb{Z}^2} f(x) \int e^{-i\lambda y} f(y - x) dy \\
&= \widehat{f}(\lambda) \sum_{x \in \mathbb{Z}^2} f(x) e^{-i\lambda x} \\
&= \widehat{f}(\lambda) \sum_{x \in \mathbb{Z}^2} \left(\frac{1}{(2\pi)^2} \int e^{ipx} \widehat{f}(p) dp \right) e^{-i\lambda x}.
\end{aligned} \quad (3.46)$$

For $x \in \mathbb{Z}^2$

$$\int e^{ipx} \widehat{f}(p) dp = \sum_{u \in \mathbb{Z}^2} \int_{[-\pi, \pi]^2} e^{ipx} \widehat{f}(p + 2\pi u) dp \quad (3.47)$$

and using Fourier inversion

$$\begin{aligned}
& \sum_{x \in \mathbb{Z}^2} \left(\int e^{ipx} \widehat{f}(p) dp \right) e^{-i\lambda x} \\
&= \sum_{u \in \mathbb{Z}^2} \sum_{x \in \mathbb{Z}^2} \left(\int_{[-\pi, \pi]^2} e^{ipx} \widehat{f}(p + 2\pi u) dp \right) e^{-i\lambda x} \\
&= (2\pi)^2 \sum_{u \in \mathbb{Z}^2} \widehat{f}(\lambda + 2\pi u).
\end{aligned} \tag{3.48}$$

Thus from (3.46) we find that

$$\int e^{-i\lambda y} f \circledast f(y) dy = \widehat{f}(\lambda) \sum_{u \in \mathbb{Z}^2} \widehat{f}(\lambda + 2\pi u). \tag{3.49}$$

Since $\widehat{f}_r(\lambda) = \widehat{f}(r\lambda)$ we see that for any $r > 0$

$$\int e^{-i\lambda y} (f_r \circledast f_r)(y) dy = \widehat{f}(r\lambda) \sum_{u \in \mathbb{Z}^2} \widehat{f}(r\lambda + 2\pi r u). \tag{3.50}$$

Then if $r \geq 1$, using the fact that $\widehat{f}(\lambda)$ is supported in $(-\pi, \pi)^2$, we obtain (3.45). \square

Taking $f \in \mathcal{S}(\mathbb{R}^2)$ with $\widehat{f}(\lambda)$ supported in $(-\pi, \pi)^2$, Lemma 3.3 will follow from Theorem 4.1 of the next section. \square

4 Intersections of Random Walks

Let $S_1(n), S_2(n)$ be independent copies of the symmetric random walk $S(n)$ in \mathbb{Z}^2 with a finite second moment.

Let f be a positive symmetric function in the Schwartz space $\mathcal{S}(\mathbb{R}^2)$ with $\int f dx = 1$ and \widehat{f} supported in $(-\pi, \pi)^2$. Given $\epsilon > 0$, and with the notation of the last section, let us define the link

$$I_{n,\epsilon}(y) = \sum_{n_1, n_2=1}^n f_{(b_n^{-1}n)^{1/2}\epsilon} \circledast f_{(b_n^{-1}n)^{1/2}\epsilon}(S_2(n_2) - S_1(n_1) + y) \tag{4.1}$$

with $I_{n,\epsilon} = I_{n,\epsilon}(0)$.

Theorem 4.1. *For any $\lambda > 0$*

$$\limsup_{\epsilon \rightarrow 0} \limsup_{n \rightarrow \infty} \sup_y \left(\frac{1}{b_n} \log \mathbb{E} \left(\exp \left\{ \lambda \left| \frac{I_n(y) - I_{n,\epsilon}(y)}{b_n^{-1}n} \right|^{1/2} \right\} \right) \right) = 0. \tag{4.2}$$

Proof of Theorem 4.1. We prove this result by obtaining moment estimates using Fourier analysis. The fact that \widehat{f} is supported in $(-\pi, \pi)^2$ plays a critical role by allowing us to express both $I_n(y)$ and $I_{n,\epsilon}(y)$ as Fourier integrals over the same region. Compare (4.4), (4.6) and (4.9).

We have

$$\begin{aligned} & \frac{1}{b_n^{-1}n} I_n(y) \\ &= \frac{1}{b_n^{-1}n} \sum_{n_1, n_2=1}^n \delta(S_1(n_1), S_2(n_2) + y) \\ &= \frac{1}{b_n^{-1}n(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{[-\pi, \pi]^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1))} dp \right] \end{aligned} \quad (4.3)$$

where from now on we work modulo $\pm\pi$. Then by scaling we have

$$\begin{aligned} & \frac{1}{b_n^{-1}n} I_n(y) \\ &= \frac{1}{(b_n^{-1}n)^2(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1)) / (b_n^{-1}n)^{1/2}} dp \right] \end{aligned} \quad (4.4)$$

As in (4.3)-(4.4), using Lemma 3.4, the fact that $\epsilon(b_n^{-1}n)^{1/2} \geq 1$ for $\epsilon > 0$ fixed and large enough n , and abbreviating $\widehat{h} = (\widehat{f})^2$

$$\begin{aligned} & \frac{1}{b_n^{-1}n} I_{n,\epsilon}(y) \\ &= \frac{1}{b_n^{-1}n(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{\mathbb{R}^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1))} \widehat{h}(\epsilon(b_n^{-1}n)^{1/2}p) dp \right] \\ &= \frac{1}{(b_n^{-1}n)^2(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{\mathbb{R}^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1)) / (b_n^{-1}n)^{1/2}} \widehat{h}(\epsilon p) dp \right]. \end{aligned} \quad (4.5)$$

Using our assumption that \widehat{h} supported in $[-\pi, \pi]^2$, and that $\epsilon^{-1} \leq (b_n^{-1}n)^{1/2}$ for $\epsilon > 0$ fixed and large enough n , we have that

$$\begin{aligned} & \frac{1}{b_n^{-1}n} I_{n,\epsilon}(y) \\ &= \frac{1}{(b_n^{-1}n)^2(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{\epsilon^{-1}[-\pi, \pi]^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1)) / (b_n^{-1}n)^{1/2}} \widehat{h}(\epsilon p) dp \right] \\ &= \frac{1}{(b_n^{-1}n)^2(2\pi)^2} \sum_{n_1, n_2=1}^n \left[\int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^2} e^{ip \cdot (S_2(n_2) + y - S_1(n_1)) / (b_n^{-1}n)^{1/2}} \widehat{h}(\epsilon p) dp \right]. \end{aligned} \quad (4.6)$$

To prove (4.2) it suffices to show that for each $\lambda > 0$ we have

$$\begin{aligned} & \sup_y \mathbb{E} \left(\exp \left\{ \lambda \left| \frac{I_n(y) - I_{n,\epsilon}(y)}{b_n^{-1}n} \right|^{1/2} \right\} \right) \\ & \leq C b_n (1 - C \lambda \epsilon^{m/4})^{-1} (1 + C \lambda \epsilon^{1/4} b_n^{1/2}) e^{C \lambda^2 \epsilon^{1/2} b_n}. \end{aligned} \quad (4.7)$$

for some $C < \infty$ and all $\epsilon > 0$ sufficiently small.

We begin by expanding

$$\begin{aligned} & \mathbb{E} \left(\exp \left\{ \lambda \left| \frac{I_n(y) - I_{n,\epsilon}(y)}{b_n^{-1}n} \right|^{1/2} \right\} \right) \\ & = \sum_{m=0}^{\infty} \frac{\lambda^m}{m!} \mathbb{E} \left(\left| \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right|^{m/2} \right) \\ & \leq \sum_{m=0}^{\infty} \frac{\lambda^m}{m!} \left(\mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^{2m} \right) \right)^{1/4} \end{aligned} \quad (4.8)$$

By (4.4), (4.6) and the symmetry of S_1 we have

$$\begin{aligned} & \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \\ & = \frac{1}{(b_n^{-1}n)^{2m} (2\pi)^{2m}} \sum_{\substack{n_{1,j}, n_{2,j}=1 \\ j=1, \dots, m}}^n \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \\ & \quad \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot (S_2(n_{2,j}) + y + S_1(n_{1,j})) / (b_n^{-1}n)^{1/2}} \right) \prod_{j=1}^m (1 - \widehat{h}(\epsilon p_j)) dp_j. \end{aligned} \quad (4.9)$$

Then

$$\begin{aligned} & \left| \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \right| \\ & \leq \frac{1}{(b_n^{-1}n)^{2m} (2\pi)^{2m}} \sum_{\substack{n_{1,j}=1 \\ j=1, \dots, m}}^n \sum_{\substack{n_{2,j}=1 \\ j=1, \dots, m}}^n \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \\ & \quad \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S_1(n_{1,j}) / (b_n^{-1}n)^{1/2}} \right) \right| \\ & \quad \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S_2(n_{2,j}) / (b_n^{-1}n)^{1/2}} \right) \right| \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j. \end{aligned} \quad (4.10)$$

By the Cauchy-Schwarz inequality

$$\int_{(b_n^{-1}n)^{1/2}[-\pi,\pi]^{2m}} \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S_1(n_{1,j}) / (b_n^{-1}n)^{1/2}} \right) \right| \quad (4.11)$$

$$\left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S_2(n_{2,j}) / (b_n^{-1}n)^{1/2}} \right) \right| \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j$$

$$\leq \prod_{i=1}^2 \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi,\pi]^{2m}} \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S(n_{i,j}) / (b_n^{-1}n)^{1/2}} \right) \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2}. \quad (4.12)$$

Thus

$$\left| \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \right|^{1/2} \quad (4.13)$$

$$\leq \sum_{\substack{n_j=1 \\ j=1,\dots,m}}^n \frac{1}{(b_n^{-1}n)^m (2\pi)^m} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi,\pi]^{2m}} \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S(n_j) / (b_n^{-1}n)^{1/2}} \right) \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2}$$

For any permutation ψ of $\{1, \dots, m\}$ let

$$D_m(\psi) = \{(n_1, \dots, n_m) \mid 1 \leq n_{\psi(1)} \leq \dots \leq n_{\psi(m)} \leq n\}. \quad (4.14)$$

Using the (non-disjoint) decomposition

$$\{1, \dots, n\}^m = \bigcup_{\psi} D_m(\psi)$$

we have from (4.13) that

$$\left| \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \right|^{1/2} \quad (4.15)$$

$$\leq \sum_{\psi} \sum_{D_m(\psi)} \frac{1}{(b_n^{-1}n)^m (2\pi)^m} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi,\pi]^{2m}} \left| \mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S(n_j) / (b_n^{-1}n)^{1/2}} \right) \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2}$$

where the first sum is over all permutations ψ of $\{1, \dots, m\}$.

Set

$$\phi(u) = \mathbb{E} \left(e^{iu \cdot S(1)} \right). \quad (4.16)$$

It follows from our assumptions that $\phi(u) \in C^2$, $\frac{\partial}{\partial u_i} \phi(0) = 0$ and $\frac{\partial^2}{\partial u_i \partial u_j} \phi(0) = -\mathbb{E} (S_{(i)}(1)S_{(j)}(1))$ where $S(1) = (S_{(1)}(1), S_{(2)}(1))$ so that for some $\delta > 0$

$$\phi(u) = 1 - \mathbb{E} ((u \cdot S(1))^2) / 2 + o(|u|^2), \quad |u| \leq \delta. \quad (4.17)$$

Then for some $c_1 > 0$

$$|\phi(u)| \leq e^{-c_1|u|^2}, \quad |u| \leq \delta. \quad (4.18)$$

Strong aperiodicity implies that $|\phi(u)| < 1$ for $u \neq 0$ and $u \in [-\pi, \pi]^2$. In particular, we can find $b < 1$ such that $|\phi(u)| \leq b$ for $\delta \leq |u|$ and $u \in [-\pi, \pi]^2$. But clearly we can choose $c_2 > 0$ so that $b \leq e^{-c_2|u|^2}$ for $u \in [-\pi, \pi]^2$. Setting $c = \min(c_1, c_2) > 0$ we then have

$$|\phi(u)| \leq e^{-c|u|^2}, \quad u \in [-\pi, \pi]^2. \quad (4.19)$$

On $D_m(\psi)$ we can write

$$\sum_{j=1}^m p_j \cdot S(n_j) = \sum_{j=1}^m \left(\sum_{i=j}^m p_{\psi(i)} (S(n_{\psi(j)}) - S(n_{\psi(j-1)})) \right). \quad (4.20)$$

Hence on $D_m(\psi)$

$$\mathbb{E} \left(e^{i \sum_{j=1}^m p_j \cdot S(n_j) / (b_n^{-1}n)^{1/2}} \right) = \prod_{j=1}^m \phi \left(\left(\sum_{i=j}^m p_{\psi(i)} \right) / (b_n^{-1}n)^{1/2} \right)^{(n_{\psi(j)} - n_{\psi(j-1)})}. \quad (4.21)$$

Now it is clear that

$$\begin{aligned} & \sum_{D_m(\psi)} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \right. \\ & \left. \left| \prod_{j=1}^m \phi \left(\left(\sum_{i=j}^m p_{\psi(i)} \right) / (b_n^{-1}n)^{1/2} \right)^{(n_{\psi(j)} - n_{\psi(j-1)})} \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2} \\ &= \sum_{1 \leq n_{\psi(1)} \leq \dots \leq n_{\psi(m)} \leq n} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \right. \\ & \left. \left| \prod_{j=1}^m \phi \left(\left(\sum_{i=j}^m p_{\psi(i)} \right) / (b_n^{-1}n)^{1/2} \right)^{(n_{\psi(j)} - n_{\psi(j-1)})} \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2} \end{aligned} \quad (4.22)$$

is independent of the permutation ψ . Hence writing

$$u_j = \sum_{i=j}^m p_i \quad (4.23)$$

we have from (4.15) that

$$\begin{aligned}
& \left| \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \right|^{1/2} \\
& \leq m! \sum_{1 \leq n_1 \leq \dots \leq n_m \leq n} \frac{1}{(b_n^{-1}n)^m (2\pi)^m} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \right. \\
& \quad \left. \left| \prod_{j=1}^m \phi(u_j / (b_n^{-1}n)^{1/2})^{(n_j - n_{j-1})} \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2}.
\end{aligned} \tag{4.24}$$

For each $A \subseteq \{2, 3, \dots, m\}$ we use $D_m(A)$ to denote the subset of $\{1 \leq n_1 \leq \dots \leq n_m \leq n\}$ for which $n_j = n_{j-1}$ if and only if $j \in A$. Then we have

$$\begin{aligned}
& \left| \mathbb{E} \left(\left\{ \frac{1}{b_n^{-1}n} (I_n(y) - I_{n,\epsilon}(y)) \right\}^m \right) \right|^{1/2} \\
& \leq m! \sum_{A \subseteq \{2, 3, \dots, m\}} \sum_{D_m(A)} \frac{1}{(b_n^{-1}n)^m (2\pi)^m} \left\{ \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \right. \\
& \quad \left. \left| \prod_{j=1}^m \phi(u_j / (b_n^{-1}n)^{1/2})^{(n_j - n_{j-1})} \right|^2 \prod_{j=1}^m |1 - \widehat{h}(\epsilon p_j)| dp_j \right\}^{1/2}.
\end{aligned} \tag{4.25}$$

For any $u \in \mathbb{R}^d$ let \widetilde{u} denote the representative of $u \bmod (b_n^{-1}n)^{1/2}2\pi\mathbb{Z}^2$ of smallest absolute value. We note that

$$|\widetilde{-u}| = |\widetilde{u}|, \quad \text{and} \quad |\widetilde{u+v}| = |\widetilde{u} + \widetilde{v}| \leq |\widetilde{u}| + |\widetilde{v}|. \tag{4.26}$$

Using the periodicity of ϕ we see that (4.19) implies that for all u

$$|\phi(u / (b_n^{-1}n)^{1/2})| \leq e^{-c|\widetilde{u}|^2 / (b_n^{-1}n)}. \tag{4.27}$$

Then we have that on $\{1 \leq n_1 \leq \dots \leq n_m \leq n\}$

$$\left| \prod_{j=1}^m \phi(u_j / (b_n^{-1}n)^{1/2})^{(n_j - n_{j-1})} \right|^2 \leq \prod_{j=1}^m e^{-c|\widetilde{u}_j|^2 (n_j - n_{j-1}) / (b_n^{-1}n)} \tag{4.28}$$

Using $|1 - \widehat{h}(\epsilon p_j)| \leq c\epsilon^{1/2}|p_j|^{1/2}$ we bound the integral in (4.25) by

$$c^m \epsilon^{m/2} \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \prod_{j=1}^m e^{-c|\widetilde{u}_j|^2 (n_j - n_{j-1}) / (b_n^{-1}n)} |p_j|^{1/2} dp_j. \tag{4.29}$$

Using (4.23) and (4.26) we have that

$$\prod_{j=1}^m |p_j|^{1/2} \leq \prod_{j=1}^m (|\widetilde{u}_j|^{1/2} + |\widetilde{u}_{j+1}|^{1/2}) \tag{4.30}$$

and when we expand the right hand side as a sum of monomials we can be sure that no factor $|\tilde{u}_k|^{1/2}$ appears more than twice. Thus we see that we can bound (4.29) by

$$C^m \epsilon^{m/2} \max_{h(j)} \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \prod_{j=1}^m e^{-c|\tilde{u}_j|^2(n_j - n_{j-1})/(b_n^{-1}n)} |\tilde{u}_j|^{h(j)/2} dp_j \quad (4.31)$$

where the max runs over the the set of functions $h(j)$ taking values 0, 1 or 2 and such that $\sum_j h(j) = m$. Here we used the fact that the number of ways to choose the $\{h(j)\}$ is bounded by the number of ways of dividing m objects into 3 groups, which is 3^m . Changing variables, we thus need to bound

$$\int_{\Lambda_n} \prod_{j=1}^m e^{-c|\tilde{u}_j|^2(n_j - n_{j-1})/(b_n^{-1}n)} |\tilde{u}_j|^{h(j)/2} du_j \quad (4.32)$$

where, see (4.23),

$$\Lambda_n = \{(u_1, \dots, u_m) \mid u_j - u_{j+1} \in (b_n^{-1}n)^{1/2}[-\pi, \pi]^2, \forall j\}. \quad (4.33)$$

Let C_n denote the rectangle $(b_n^{-1}n)^{1/2}[-\pi, \pi]^2$ and let us call any rectangle of the form $2\pi k + C_n$, where $k \in \mathbb{Z}^2$, an elementary rectangle. Note that any rectangle of the form $v + C_n$, where $v \in \mathbb{R}^2$, can be covered by 4 elementary rectangles. Hence for any $v \in \mathbb{R}^2$ and $1 \leq s \leq n$

$$\begin{aligned} & \int_{v+C_n} e^{-c\frac{s}{b_n^{-1}n}|\tilde{u}|^2} |\tilde{u}|^{h/2} du \\ & \leq 4 \int_{\mathbb{R}^2} e^{-c\frac{s}{(b_n^{-1}n)}|u|^2} |u|^{h/2} du \\ & \leq C \left(\frac{s}{b_n^{-1}n} \right)^{-(1+h/4)}. \end{aligned} \quad (4.34)$$

Similarly

$$\int_{v+C_n} |\tilde{u}|^{h/2} du \leq C(b_n^{-1}n)^{(1+h/4)}. \quad (4.35)$$

We now bound (4.32) by bounding successively the integration with respect to u_1, \dots, u_m . Consider first the du_1 integral, fixing u_2, \dots, u_m . By (4.33) the du_1 integral is over the rectangle $u_2 + C_n$, hence the factors involving u_1 can be bounded using (4.34). Proceeding inductively, using (4.33) when $n_j - n_{j-1} > 0$ and (4.35) when $n_j = n_{j-1}$, leads to the following bound of (4.32), and hence of (4.29) on $D_m(A)$:

$$\begin{aligned} & c^m \epsilon^{m/2} \int_{(b_n^{-1}n)^{1/2}[-\pi, \pi]^{2m}} \prod_{j=1}^m e^{-c|\tilde{u}_j|^2(n_j - n_{j-1})/(b_n^{-1}n)} |p_j|^{1/2} dp_j \\ & \leq C^m \epsilon^{m/2} \prod_{j \in A} (b_n^{-1}n)^{(1+h(j)/4)} \prod_{j \in A^c} \left(\frac{(n_j - n_{j-1})}{b_n^{-1}n} \right)^{-(1+h(j)/4)}. \end{aligned} \quad (4.36)$$

Here A^c means the complement of A in $\{1, \dots, m\}$, so that A^c always contains 1. If $A^c = \{i_1, \dots, i_k\}$ where $i_1 < \dots < i_k$ we then obtain for the sum in (4.25) over $D_m(A)$, the bound

$$C^m \epsilon^{m/4} \max_{h(j)} \frac{1}{(b_n^{-1}n)^m} \prod_{j \in A} (b_n^{-1}n)^{(1+h(j)/4)} \quad (4.37)$$

$$\sum_{1 \leq n_{i_1} < \dots < n_{i_k} \leq n} \prod_{j \in A^c} \left(\frac{(n_j - n_{j-1})}{b_n^{-1}n} \right)^{-(1+h(j)/4)/2}$$

Note that

$$(b_n^{-1}n)^{(1+h(j)/4)/2} \frac{1}{b_n^{-1}n} \rightarrow 0 \text{ as } n \rightarrow \infty. \quad (4.38)$$

Using this to bound the product over $j \in A$, and then bounding the sum by an integral, we can bound (4.37) by

$$C^m \epsilon^{m/4} \max_{h(j)} \sum_{1 \leq n_{i_1} < \dots < n_{i_k} \leq n} \prod_{j \in A^c} \left(\frac{(n_j - n_{j-1})}{b_n^{-1}n} \right)^{-(1+h(j)/4)/2} \frac{1}{b_n^{-1}n} \quad (4.39)$$

$$\leq C^m \epsilon^{m/4} \max_{h(j)} \int_{0 \leq r_{i_1} < \dots < r_{i_k} \leq b_n} \prod_{j \in A^c} (r_j - r_{j-1})^{-(1/2+h(j)/8)} dr_j$$

$$\leq C^m \epsilon^{m/4} \max_{h(j)} \frac{b_n^{\sum_{j \in A^c} (1/2 - h(j)/8)}}{\Gamma(\sum_{j \in A^c} (1/2 - h(j)/8))}$$

Using this together with (4.25), but with m replaced by $2m$, and the fact that $(2m!)^{1/2}/m! \leq 2^m$, we see that (4.7) is bounded by

$$\sum_{m=0}^{\infty} C^m \lambda^m \epsilon^{m/4} \left(\sum_{A \subseteq \{2,3,\dots,2m\}} \max_{h(j)} \frac{b_n^{\sum_{j \in A^c} (1/2 - h(j)/8)}}{\Gamma(\sum_{j \in A^c} (1/2 - h(j)/8))} \right)^{1/2}. \quad (4.40)$$

We have $\sum_{A \subseteq \{1,2,3,\dots,2m\}} 1 = 2^{2m}$. Then noting that $\sum_{j \in A^c} (1/2 - h(j)/8)$ is an integer multiple of $1/8$ which is always less than m , we can bound the last line by

$$\sum_{l=0}^{\infty} \left(\sum_{m=l}^{\infty} C^m \lambda^m \epsilon^{m/4} \right) \sum_{j=0}^7 \left(\frac{b_n^{l+j/8}}{\Gamma(l+j/8)} \right)^{1/2} \quad (4.41)$$

$$\leq C b_n \sum_{l=0}^{\infty} \left(\sum_{m=l}^{\infty} C^m \lambda^m \epsilon^{m/4} \right) \left(\frac{b_n^l}{\Gamma(l)} \right)^{1/2}$$

$$\leq C b_n (1 - C \lambda \epsilon^{m/4})^{-1} \sum_{l=0}^{\infty} C^l \lambda^l |\epsilon|^{l/4} b_n^{l/2} \left(\frac{1}{\Gamma(l)} \right)^{1/2}$$

for $\epsilon > 0$ sufficiently small.

(4.7) then follows from the fact that for any $a > 0$

$$\begin{aligned}
& \sum_{l=0}^{\infty} a^l \left(\frac{1}{\Gamma(l)} \right)^{1/2} \\
&= \sum_{m=0}^{\infty} \left(a^{2m} \left(\frac{1}{\Gamma(2m)} \right)^{1/2} + a^{2m+1} \left(\frac{1}{\Gamma(2m+1)} \right)^{1/2} \right) \\
&\leq C(1+a) \sum_{m=0}^{\infty} a^{2m} \left(\frac{1}{\Gamma(2m)} \right)^{1/2} \\
&\leq C(1+a)e^{Ca^2}.
\end{aligned} \tag{4.42}$$

□

Remark 4.2. It follows from the proof that in fact for $\rho > 0$ sufficiently small, for any $\lambda > 0$

$$\begin{aligned}
& \limsup_{\epsilon \rightarrow 0} \limsup_{n \rightarrow \infty} \sup_y \\
& \frac{1}{b_n} \log \mathbb{E} \left(\exp \left\{ \lambda \left| \frac{I_n(y) - I_{n,\epsilon}(y)}{\epsilon^\rho b_n^{-1} n} \right|^{1/2} \right\} \right) = 0.
\end{aligned} \tag{4.43}$$

5 Theorem 1.2: Upper bound for $\mathbb{E} B_n - B_n$

Proof of Theorem 1.2.

We prove (1.10):

$$\begin{aligned}
-C_1 &\leq \liminf_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\
&\leq \limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \leq -C_2
\end{aligned} \tag{5.1}$$

for any $\{b_n\}$ satisfying (1.8).

In this section we prove the upper bound for (5.1). We will derive this by using an analogous bound for the renormalized self-intersection local time of planar Brownian motion. Let $t > 0$ and write $K = \lceil t^{-1} b_n \rceil$. Divide $[1, n]$ into $K > 1$ disjoint subintervals $(n_0, n_1], \dots, (n_{K-1}, n_K]$, each of length $\lfloor n/K \rfloor$ or $\lfloor n/K \rfloor + 1$. Notice that

$$\begin{aligned}
\mathbb{E} B_n - B_n &\leq \sum_{i=1}^K \left[\mathbb{E} B((n_{i-1}, n_i]_{<}^2) - B((n_{i-1}, n_i]_{<}^2) \right] \\
&\quad + \mathbb{E} B_n - \sum_{i=1}^K \mathbb{E} B((n_{i-1}, n_i]_{<}^2)
\end{aligned} \tag{5.2}$$

By (2.39),

$$\begin{aligned}
\sum_{i=1}^K \mathbb{E} B((n_{i-1}, n_i]_{<}^2) &= \sum_{i=1}^K \mathbb{E} B_{n_i - n_{i-1}} \\
&= \sum_{i=1}^K \left[\frac{1}{(2\pi)\sqrt{\det \Gamma}} (n/K) \log(n/K) + O(n/K) \right] \\
&= \frac{1}{(2\pi)\sqrt{\det \Gamma}} n \log(n/K) + O(n)
\end{aligned} \tag{5.3}$$

With $K > 1$, the error term can be taken to be independent of t and $\{b_n\}$. Thus, by (2.39), there is constant $\log a > 0$ independent of t and $\{b_n\}$ such that

$$\begin{aligned}
\mathbb{E} B_n - \sum_{j=1}^K \mathbb{E} B((n_{i-1}, n_i]_{<}^2) \\
\leq \frac{1}{(2\pi)\sqrt{\det \Gamma}} n \left(\log(t^{-1} b_n) + \log a \right).
\end{aligned} \tag{5.4}$$

It is here that we use the condition that $\mathbb{E} |S_1|^{2+\delta} < \infty$ for some $\delta > 0$, needed for (2.39).

By first using Chebyshev's inequality, then using (5.2), (5.4) and the independence of the $B((n_{i-1}, n_i]_{<}^2)$, for any $\phi > 0$,

$$\begin{aligned}
&\mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\
&\leq \exp \left\{ -\phi b_n \log b_n \right\} \mathbb{E} \exp \left\{ -2\pi\phi\sqrt{\det \Gamma} \frac{b_n}{n} (B_n - \mathbb{E} B_n) \right\} \\
&\leq \exp \left\{ \phi b_n (\log a - \log t) \right\} \left(\mathbb{E} \exp \left\{ -2\pi\phi\sqrt{\det \Gamma} \frac{b_n}{n} (B_{[n/K]} - \mathbb{E} B_{[n/K]}) \right\} \right)^K
\end{aligned} \tag{5.5}$$

By (16, Theorem 1.2),

$$\sqrt{\det \Gamma} \frac{b_n}{n} (B_{[n/K]} - \mathbb{E} B_{[n/K]}) \xrightarrow{d} \tilde{\gamma}_t, \quad (n \rightarrow \infty) \tag{5.6}$$

where $\tilde{\gamma}_t$ is the renormalized self-intersection local time of planar Brownian motion $\{W_s\}$ up to time t . By Lemma 2.5 and the dominated convergence theorem,

$$\mathbb{E} \exp \left\{ -2\pi\phi\sqrt{\det \Gamma} \frac{b_n}{n} (B_{[n/K]} - \mathbb{E} B_{[n/K]}) \right\} \longrightarrow \mathbb{E} \exp \left\{ -2\pi\phi t \tilde{\gamma}_1 \right\}, \quad (n \rightarrow \infty) \tag{5.7}$$

where we used the scaling $\tilde{\gamma}_t \stackrel{d}{=} t \tilde{\gamma}_1$.

Thus,

$$\begin{aligned}
&\limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\
&\leq \phi (\log a - \log t) + \frac{1}{t} \log \mathbb{E} \exp \left\{ -2\pi\phi t \tilde{\gamma}_1 \right\} \\
&= \phi \log(a\phi) + \frac{1}{t} \log \mathbb{E} \exp \left\{ -(\phi t) \log(\theta t) - 2\pi(\phi t) \tilde{\gamma}_1 \right\}
\end{aligned} \tag{5.8}$$

By (2, p. 3233), the limit

$$C \equiv \lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{E} \exp \left\{ -t \log t - 2\pi t \tilde{\gamma}_1 \right\} \quad (5.9)$$

exists. Hence

$$\begin{aligned} \limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\ \leq \phi \log(a\phi) + C\phi. \end{aligned} \quad (5.10)$$

Taking the minimizer $\phi = a^{-1}e^{-(1+C)}$ we have

$$\begin{aligned} \limsup_{n \rightarrow \infty} b_n^{-1} \log \mathbb{P} \left\{ \mathbb{E} B_n - B_n \geq (2\pi)^{-1} \det(\Gamma)^{-1/2} n \log b_n \right\} \\ \leq -a^{-1}e^{-(1+C)}. \end{aligned} \quad (5.11)$$

This proves the upper bound for (5.1).

6 Theorem 1.2: Lower bound for $\mathbb{E} B_n - B_n$

In this section we complete the proof of Theorem 1.2 by proving the lower bound for (5.1). As before, we will see that the contribution to B_n in the scale of interest to us from i, j near the diagonal is almost deterministic and therefore will be comparable to $\mathbb{E} B_n$. The heart of the proof of our Theorem is to show that with probability which is not too small the contributions to B_n from i, j far from the diagonal is not too large. We accomplish this showing that with probability which is not too small we can assure that the random walk has a drift so that $S_i \neq S_j$ for i, j far apart.

Let $\mathcal{F}_k = \sigma\{X_i : i \leq k\}$. Let us assume for simplicity that the covariance matrix for the random walk is the identity; routine modifications are all that are needed for the general case. We write Θ for $(2\pi)^{-1} \det(\Gamma)^{-1/2} = (2\pi)^{-1}$. We write $D(x, r)$ for the disc of radius r in \mathbb{Z}^2 centered at x .

Let $K = [b_n]$ and $L = n/K$. Let us divide $\{1, 2, \dots, n\}$ into K disjoint contiguous blocks, each of length strictly between $L/2$ and $3L/2$. Denote the blocks J_1, \dots, J_K . Let $v_i = \#(J_i)$, $w_i = \sum_{j=1}^i v_j$. Let

$$B_{v_i}^{(i)} = \sum_{j, k \in J_i, j < k} \delta(S_j, S_k), \quad A_i = \sum_{j \in J_{i-1}, k \in J_i} \delta(S_j, S_k). \quad (6.1)$$

Define the following sets:

$$\begin{aligned} F_{i,1} &= \{S_{w_i} \in D(i\sqrt{L}, \sqrt{L}/16)\}, \\ F_{i,2} &= \{S(J_i) \subset [(i-1)\sqrt{L} - \sqrt{L}/8, i\sqrt{L} + \sqrt{L}/8] \times [-\sqrt{L}/8, \sqrt{L}/8]\}, \\ F_{i,3} &= \{B_{v_i}^{(i)} - \mathbb{E} B_{v_i}^{(i)} \leq \kappa_1 L\}, \\ F_{i,4} &= \left\{ \sum_{j \in J_i} 1_{D(x, r\sqrt{L})}(S_j) \leq \kappa_2 r L \text{ for all } x \in D(i\sqrt{L}, 3\sqrt{L}), 1/\sqrt{L} < r < 2 \right\}, \\ F_{i,5} &= \{A_i < \kappa_3 L\}, \end{aligned}$$

where $\kappa_1, \kappa_2, \kappa_3$ are constants that will be chosen later and do not depend on K or L . Let

$$C_i = F_{i,1} \cap F_{i,2} \cap F_{i,3} \cap F_{i,4} \cap F_{i,5} \quad (6.2)$$

and

$$E = \bigcap_{i=1}^K C_i. \quad (6.3)$$

We want to show

$$\mathbb{P}(C_i \mid \mathcal{F}_{w_{i-1}}) \geq c_1 > 0 \quad (6.4)$$

on the event $C_1 \cap \dots \cap C_{i-1}$. Once we have (6.4), then

$$\mathbb{P}(\bigcap_{i=1}^m C_i) = \mathbb{E} \left(\mathbb{P}(C_m \mid \mathcal{F}_{w_{m-1}}); \bigcap_{i=1}^{m-1} C_i \right) \geq c_1 \mathbb{P}(\bigcap_{i=1}^{m-1} C_i), \quad (6.5)$$

and by induction

$$\mathbb{P}(E) = \mathbb{P}(\bigcap_{i=1}^K C_i) \geq c_1^K = e^{K \log c_1} = e^{-c_2 K}. \quad (6.6)$$

On the set E , we see that $S(J_i) \cap S(J_j) = \emptyset$ if $|i - j| > 1$. So we can write

$$B_n = \sum_{k=1}^K (B_{v_k}^{(k)} - \mathbb{E} B_{v_k}^{(k)}) + \sum_{k=1}^K \mathbb{E} B_{v_k}^{(k)} + \sum_{k=1}^K A_k. \quad (6.7)$$

On the event E , each $B_{v_k}^{(k)} - \mathbb{E} B_{v_k}^{(k)}$ is bounded by $\kappa_1 L$ and each A_k is bounded by $\kappa_3 L$. By (2.38), each $\mathbb{E} B_{v_k}^{(k)} = \Theta v_k \log v_k + O(L) = \Theta v_k \log L + O(v_k)$. Therefore

$$B_n \leq \kappa_1 K L + \Theta K L \log L + O(n) + \kappa_3 K L, \quad (6.8)$$

and using (2.38) again,

$$\begin{aligned} \mathbb{E} B_n - B_n &\geq \Theta n \log n - c_3 n - \Theta n \log(n/b_n) \\ &= \Theta n \log b_n - c_3 n \end{aligned} \quad (6.9)$$

on the event E . We conclude that

$$\mathbb{P}(\mathbb{E} B_n - B_n \geq \Theta n \log b_n - c_3 n) \geq e^{-c_2 b_n}. \quad (6.10)$$

We apply (6.10) with b_n replaced by $b'_n = c_4 b_n$, where $\Theta \log c_4 = c_3$. Then

$$\Theta n \log b'_n - c_3 n = \Theta n \log b_n + \Theta n \log c_4 - c_3 n = \Theta n \log b_n. \quad (6.11)$$

We then obtain

$$\mathbb{P}(\mathbb{E} B_n - B_n \geq \Theta n \log b_n) = \mathbb{P}(\mathbb{E} B_n - B_n \geq \Theta n \log b'_n - c_3 n) \geq e^{-c_2 b'_n}, \quad (6.12)$$

which would complete the proof of the lower bound for (5.1), hence of Theorem 1.2.

So we need to prove (6.4). By scaling and the support theorem for Brownian motion (see (1, Theorem I.6.6)), if W_t is a planar Brownian motion and $|x| \leq \sqrt{L}/16$, then

$$\begin{aligned} \mathbb{P}^x \left(W_{v_i} \in D(\sqrt{L}, \sqrt{L}/16) \text{ and} \right. \\ \left. \{W_s; 0 \leq s \leq v_i\} \subset [-\sqrt{L}/8, 9\sqrt{L}/8] \times [-\sqrt{L}/8, \sqrt{L}/8] \right) > c_5, \end{aligned} \quad (6.13)$$

where c_5 does not depend on L . Using Donsker's invariance principle for random walks with finite second moments together with the Markov property,

$$\mathbb{P}(F_{i,1} \cap F_{i,2} \mid F_{w_{i-1}}) > c_6. \quad (6.14)$$

By Lemma 2.3, for $L/2 \leq \ell \leq 3L/2$

$$\mathbb{P}(B_\ell - \mathbb{E} B_\ell > \kappa_1 L) \leq c_6/2 \quad (6.15)$$

if we choose κ_1 large enough. Again using the Markov property,

$$\mathbb{P}(F_{i,1} \cap F_{i,2} \cap F_{i,3} \mid F_{w_{i-1}}) > c_6/2. \quad (6.16)$$

Now let us look at $F_{i,4}$. By (17, p. 75), $\mathbb{P}(S_j = y) \leq c_7/j$ with c_7 independent of $y \in \mathbb{Z}^2$ so that

$$\mathbb{P}(S_j \in D(x, r\sqrt{L})) = \sum_{y \in D(x, r\sqrt{L})} \mathbb{P}(S_j = y) \leq \frac{c_8 r^2 L}{j}. \quad (6.17)$$

Therefore

$$\begin{aligned} \mathbb{E} \sum_{j \in J_1} 1_{D(x, r\sqrt{L})}(S_j) &\leq \sum_{j=1}^{[2L]} \mathbb{P}(S_j \in D(x, r\sqrt{L})) \\ &\leq r^2 L + \sum_{j=r^2 L}^{[2L]} \frac{c_9 r^2 L}{j} \\ &\leq r^2 L + c_{10} L r^2 \log(1/r) \leq c_{11} L r^2 \log(1/r) \end{aligned} \quad (6.18)$$

if $1/\sqrt{L} \leq r \leq 2$. Let $C_m = \sum_{j < m} 1_{D(x, r\sqrt{L})}(S_j)$ for $m \leq [2L] + 1$ and let $C_m = C_{[2L]+1}$ for $m > L$. By the Markov property and independence,

$$\begin{aligned} \mathbb{E}[C_\infty - C_m \mid \mathcal{F}_m] &\leq 1 + \mathbb{E}[C_\infty - C_{m+1} \mid \mathcal{F}_m] \\ &\leq 1 + \mathbb{E}^{S_m} C_\infty \leq c_{12} L r^2 \log(1/r). \end{aligned} \quad (6.19)$$

By (1, Theorem I.6.11), we have

$$\mathbb{E} \exp\left(c_{13} \frac{C_{[2L]+1}}{c_{12} L r^2 \log(1/r)}\right) \leq c_{14} \quad (6.20)$$

with c_{13}, c_{14} independent of L or r . We conclude

$$\mathbb{P}\left(\sum_{j \in J_1} 1_{D(x, r\sqrt{L})}(S_j) > c_{15} L r^2 \log(1/r)\right) \leq c_{16} e^{-c_{17} c_{15}}. \quad (6.21)$$

Suppose $2^{-s} \leq r < 2^{-s+1}$ for some $s \geq 0$. If $x \in D(0, 3\sqrt{L})$, then each point in the disc $D(x, r\sqrt{L})$ will be contained in $D(x_i, 2^{-s+3}\sqrt{L})$ for some x_i , where each coordinate of x_i is an integer multiple of $2^{-s-2}\sqrt{L}$. There are at most $c_{18} 2^{2s}$ such balls, and $L r^2 \log(1/r) \leq c_{19} 2^{s/2} L r$, so

$$\mathbb{P}\left(\sup_{x \in D(0, 3\sqrt{L}), 2^{-s} \leq r < 2^{-s+1}} \sum_{j \in J_1} 1_{D(x, r\sqrt{L})}(S_j) > c_{20} r L\right) \leq c_{21} 2^{2s} e^{-c_{22} c_{20} 2^{s/2}}. \quad (6.22)$$

If we now sum over positive integers s and take $\kappa_2 = c_{20}$ large enough, we see that

$$\mathbb{P}(F_{1,4}^c) \leq c_6/4. \quad (6.23)$$

By the Markov property, we then obtain

$$\mathbb{P}(F_{i,1} \cap F_{i,2} \cap F_{i,3} \cap F_{i,4} \mid \mathcal{F}_{w_{i-1}}) > c_6/4. \quad (6.24)$$

Finally, we examine $F_{i,5}$. We will show

$$\mathbb{P}(F_{i,5}^c \mid \mathcal{F}_{w_{i-1}}) \leq c_6/8 \quad (6.25)$$

on the set $\cap_{j=1}^{i-1} C_j$ if we take κ_3 large enough. By the Markov property, it suffices to show

$$\mathbb{P}\left(\sum_{j=1}^{[2L]} \mathbf{1}_{(S_j \in G)} \geq \kappa_3 L\right) \leq c_6/8 \quad (6.26)$$

whenever $G \in \mathbb{Z}^2$ is a fixed nonrandom set consisting of $[2L]$ points satisfying the property that

$$\#(G \cap D(x, r\sqrt{L})) \leq \kappa_2 r L, \quad x \in D(0, 3\sqrt{L}), \quad 1/\sqrt{L} \leq r \leq 2. \quad (6.27)$$

We compute the expectation of

$$\sum_{j=1}^{[2L]} \mathbf{1}_{(S_j \in G \cap (D(0, 2^{-k}\sqrt{L}) \setminus D(0, 2^{-k+1}\sqrt{L})))}. \quad (6.28)$$

When $j \leq 2^{-2k}L$, then the fact that the random walk has finite second moments implies that the probability that $|S_j|$ exceeds $2^{-k+1}\sqrt{L}$ is bounded by $c_{23}j/(2^{-2k+2}L)$. When $j > 2^{-2k}L$, we use (17, p. 75), and obtain

$$\mathbb{P}(S_j \in G \cap (D(0, 2^{-k}\sqrt{L})) \leq c_{24} \frac{\kappa_2 2^{-k}L}{j}. \quad (6.29)$$

So

$$\begin{aligned} & \mathbb{E} \sum_{j=1}^{[2L]} \mathbf{1}_G(S_j) \\ & \leq \sum_k \sum_{[2L] \geq j > 2^{-2k}L} c_{24} \frac{\kappa_2 2^{-k}L}{j} + \sum_k \sum_{j \leq 2^{-2k}L} c_{23} \frac{j}{2^{-2k+2}L} \\ & \leq \sum_k (c_{25} \kappa_2 k 2^{-k}L + c_{26} 2^{-2k}L) \leq c_{27}L. \end{aligned} \quad (6.30)$$

So if take κ_3 large enough, we obtain (6.26).

This completes the proof of (6.4), hence of Theorem 1.2. \square

7 Laws of the iterated logarithm

7.1 Proof of the LIL for $B_n - \mathbb{E} B_n$

First, let S_j, S'_j be two independent copies of our random walk. Let

$$\ell(n, x) = \sum_{i=1}^n \delta(S_i, x), \quad \ell'(n, x) = \sum_{i=1}^n \delta(S'_i, x) \quad (7.1)$$

and note that

$$I_{k,n} = \sum_{i=1}^k \sum_{j=1}^n \delta(S_i, S'_j) = \sum_{x \in \mathbb{Z}^2} \ell(k, x) \ell'(n, x). \quad (7.2)$$

Lemma 7.1. *There exist constants c_1, c_2 such that*

$$\mathbb{P}(I_{k,n} > \lambda \sqrt{kn}) \leq c_1 e^{-c_2 \lambda}. \quad (7.3)$$

Proof. Clearly

$$(I_{k,n})^m = \sum_{x_1 \in \mathbb{Z}^2} \cdots \sum_{x_m \in \mathbb{Z}^2} \left(\prod_{i=1}^m \ell(k, x_i) \right) \left(\prod_{i=1}^m \ell'(n, x_i) \right) \quad (7.4)$$

Using the independence of S and S' ,

$$\mathbb{E} ((I_{k,n})^m) = \sum_{x_1 \in \mathbb{Z}^2} \cdots \sum_{x_m \in \mathbb{Z}^2} \mathbb{E} \left(\prod_{i=1}^m \ell(k, x_i) \right) \mathbb{E} \left(\prod_{i=1}^m \ell'(n, x_i) \right). \quad (7.5)$$

By Cauchy-Schwarz, this is less than

$$\begin{aligned} & \left[\sum_{x_1 \in \mathbb{Z}^2} \cdots \sum_{x_m \in \mathbb{Z}^2} \left(\mathbb{E} \left(\prod_{i=1}^m \ell(k, x_i) \right) \right)^2 \right]^{1/2} \\ & \left[\sum_{x_1 \in \mathbb{Z}^2} \cdots \sum_{x_m \in \mathbb{Z}^2} \left(\mathbb{E} \left(\prod_{i=1}^m \ell'(n, x_i) \right) \right)^2 \right]^{1/2} \\ & =: J_1^{1/2} J_2^{1/2}. \end{aligned} \quad (7.6)$$

We can rewrite

$$J_1 = \sum_{x_1 \in \mathbb{Z}^2} \cdots \sum_{x_m \in \mathbb{Z}^2} \mathbb{E} \left(\prod_{i=1}^m \ell(k, x_i) \right) \mathbb{E} \left(\prod_{i=1}^m \ell'(k, x_i) \right) = \mathbb{E} ((I_k)^m), \quad (7.7)$$

and similarly $J_2 = \mathbb{E} ((I_n)^m)$.

Therefore,

$$\begin{aligned}
& \mathbb{E} \exp(aI_{k,n}/\sqrt{kn}) \\
&= \sum_{m=0}^{\infty} \frac{a^m}{k^{m/2}n^{m/2}m!} \mathbb{E} ((I_{k,n})^m) \\
&\leq \sum_m \frac{a^m}{k^{m/2}n^{m/2}m!} (\mathbb{E} ((I_k)^m))^{1/2} (\mathbb{E} ((I_n)^m))^{1/2} \\
&\leq \left(\sum \frac{a^m}{m!} \mathbb{E} \left(\frac{I_k}{k} \right)^m \right)^{1/2} \left(\sum \frac{a^m}{m!} \mathbb{E} \left(\frac{I_n}{n} \right)^m \right)^{1/2} \\
&\leq \left(\mathbb{E} e^{aI_k/k} \right)^{1/2} \left(\mathbb{E} e^{aI_n/n} \right)^{1/2}.
\end{aligned} \tag{7.8}$$

By Lemma 2.2 this can be bounded independently of k and n if a is taken small, and our result follows. \square

We are now ready to prove the upper bound for the LIL for $B_n - \mathbb{E} B_n$. Write Ξ for $\sqrt{\det \Gamma} \kappa(2,2)^{-4}$. Recall that for any integrable random variable Z we let \bar{Z} denote $Z - \mathbb{E} Z$. Let $\varepsilon > 0$ and let $q > 1$ be chosen later. Our first goal is to get an upper bound on

$$\mathbb{P} \left(\max_{n/2 \leq k \leq n} \bar{B}_k > (1 + \varepsilon) \Xi^{-1} n \log \log n \right).$$

Let $m_0 = 2^N$, where N will be chosen later to depend only on ε and n . Let \mathcal{A}_0 be the integers of the form $n - km_0$ that are contained in $\{n/4, \dots, n\}$. For each i let \mathcal{A}_i be the set of integers of the form $n - km_0 2^{-i}$ that are contained in $\{n/4, \dots, n\}$. Given an integer k , let k_j be the largest element of \mathcal{A}_j that is less than or equal to k . For any $k \in \{n/2, \dots, n\}$, we can write

$$\bar{B}_k = \bar{B}_{k_0} + (\bar{B}_{k_1} - \bar{B}_{k_0}) + \dots + (\bar{B}_{k_N} - \bar{B}_{k_{N-1}}). \tag{7.9}$$

If $\bar{B}_k \geq (1 + \varepsilon) \Xi^{-1} n \log \log n$ for some $n/2 \leq k \leq n$, then either

- (a) $\bar{B}_{k_0} \geq (1 + \frac{\varepsilon}{2}) \Xi^{-1} n \log \log n$ for some $k_0 \in \mathcal{A}_0$; or else
- (b) for some $i \geq 1$ and some pair of consecutive elements $k_i, k'_i \in \mathcal{A}_i$, we have

$$\bar{B}_{k'_i} - \bar{B}_{k_i} \geq \frac{\varepsilon}{40i^2} \Xi^{-1} n \log \log n. \tag{7.10}$$

For each k_0 , using Theorem 1.1 and the fact that $k_0 \geq n/4$, the probability in (a) is bounded by

$$\exp(-(1 + \frac{\varepsilon}{4}) \log \log k_0) \leq c_1 (\log n)^{-(1 + \frac{\varepsilon}{4})}. \tag{7.11}$$

There are at most n/m_0 elements of \mathcal{A}_0 , so the probability in (a) is bounded by

$$\frac{n}{m_0} \frac{c_1}{(\log n)^{1 + \frac{\varepsilon}{4}}}. \tag{7.12}$$

Now let us examine the probability in (b). Fix i for the moment. Any two consecutive elements of \mathcal{A}_i are $2^{-i}m_0$ apart. Recalling the notation (2.16) we can write

$$\bar{B}_k - \bar{B}_j = \bar{B}([j + 1, k]_{<}^2) + \bar{B}([1, j] \times [j + 1, k]), \tag{7.13}$$

So

$$\begin{aligned} \mathbb{P}(\overline{B}_k - \overline{B}_j \geq \frac{\varepsilon}{40i^2} \Xi^{-1} n \log \log n) &\leq \mathbb{P}(\overline{B}([j+1, k]_{<}^2) \geq \frac{\varepsilon}{80i^2} \Xi^{-1} n \log \log n) \\ &+ \mathbb{P}(B([1, j] \times [j+1, k]) \geq \frac{\varepsilon}{80i^2} \Xi^{-1} n \log \log n). \end{aligned} \quad (7.14)$$

We bound the first term on the right by Lemma 2.3, and get the bound

$$\exp\left(-\frac{c\varepsilon}{80i^2} \frac{n \log \log n}{2^{-i} m_0}\right) \leq \exp\left(-\frac{c\varepsilon}{80i^2} 2^i (n/m_0) \log \log n\right) \quad (7.15)$$

if j and k are consecutive elements of \mathcal{A}_i . Note that $B([1, j] \times [j+1, k])$ is equal in law to $I_{j-1, k-j}$. Using Lemma 7.1, we bound the second term on the right hand side of (7.14) by

$$\begin{aligned} &c_1 \exp\left(-c_2 \frac{\varepsilon}{80i^2} \frac{n \log \log n}{\sqrt{2^{-i} m_0} \sqrt{j}}\right) \\ &\leq c_1 \exp\left(-c_2 \frac{\varepsilon}{80i^2} 2^{i/2} (n/m_0)^{1/2} \log \log n\right). \end{aligned} \quad (7.16)$$

The number of pairs of consecutive elements of \mathcal{A}_i is less than $2^{i+1}(n/m_0)$. So if we add (7.15) and (7.16) and multiply by the number of pairs, the probability of (b) occurring for a fixed i is bounded by

$$c_3 \frac{n}{m_0} 2^i \exp\left(-c_4 2^{i/2} (n/m_0)^{1/2} \log \log n / (80i^2)\right). \quad (7.17)$$

If we now sum over $i \geq 1$, we bound the probability in (b) by

$$c_5 \frac{n}{m_0} \exp\left(-c_6 (n/m_0)^{1/2} \log \log n\right). \quad (7.18)$$

We now choose m_0 to be the largest power of 2 so that $c_6(n/m_0)^{1/2} > 2$; recall n is big.

Let us use this value of m_0 and combine (7.12) and (7.18). Let $n_\ell = q^\ell$ and

$$C_\ell = \left\{ \max_{n_{\ell-1} \leq k \leq n_\ell} \overline{B}_k \geq (1+\varepsilon) \Xi^{-1} n_\ell \log \log n_\ell \right\}. \quad (7.19)$$

By our estimates, $\mathbb{P}(C_\ell)$ is summable, so for ℓ large, by Borel-Cantelli we have

$$\max_{n_{\ell-1} \leq k \leq n_\ell} \overline{B}_k \leq (1+\varepsilon) \Xi^{-1} n_\ell \log \log n_\ell. \quad (7.20)$$

By taking q sufficiently close to 1, this implies that for k large we have $\overline{B}_k \leq (1+2\varepsilon) \Xi^{-1} k \log \log k$. Since ε is arbitrary, we have our upper bound. \square

The lower bound for the first LIL is easier. Let $\delta > 0$ be small and let $n_\ell = \lceil e^{\ell^{1+\delta}} \rceil$. Let

$$D_\ell = \left\{ \overline{B}([n_{\ell-1} + 1, n_\ell]_{<}^2) \geq (1-\varepsilon) \Xi^{-1} n_\ell \log \log n_\ell \right\}. \quad (7.21)$$

Using Theorem 1.1, and the fact that $n_\ell/(n_\ell - n_{\ell-1})$ is of order 1, we see that $\sum_\ell \mathbb{P}(D_\ell) = \infty$ if $\delta < \varepsilon/(1-\varepsilon)$. The D_ℓ are independent, so by Borel-Cantelli

$$\overline{B}([n_{\ell-1} + 1, n_\ell]_{<}^2) \geq (1-\varepsilon) \Xi^{-1} n_\ell \log \log n_\ell \quad (7.22)$$

infinitely often with probability one. Note that as in (7.13) we can write

$$\overline{B}_{n_\ell} = \overline{B}([n_{\ell-1} + 1, n_\ell]_{<}^2) + \overline{B}_{n_{\ell-1}} + \overline{B}([1, n_{\ell-1}] \times [n_{\ell-1} + 1, n_\ell]). \quad (7.23)$$

By the upper bound,

$$\limsup_{\ell \rightarrow \infty} \frac{\overline{B}_{n_{\ell-1}}}{n_{\ell-1} \log \log n_{\ell-1}} \leq \Xi^{-1}$$

almost surely, which implies

$$\limsup_{\ell \rightarrow \infty} \frac{\overline{B}_{n_{\ell-1}}}{n_\ell \log \log n_\ell} = 0. \quad (7.24)$$

Since $B([1, n_{\ell-1}] \times [n_{\ell-1} + 1, n_\ell]) \geq 0$ and by (2.5)

$$\mathbb{E} B([1, n_{\ell-1}] \times [n_{\ell-1} + 1, n_\ell]) \leq c_1 \sqrt{n_{\ell-1}} \sqrt{n_\ell - n_{\ell-1}} = o(n_\ell \log \log n_\ell), \quad (7.25)$$

using (7.22)-(7.25) yields the lower bound. \square

7.2 LIL for $\mathbb{E} B_n - B_n$

Let $\Delta = 2\pi\sqrt{\det \Gamma}$. Let us write $J_n = \mathbb{E} B_n - B_n$.

First we do the upper bound. Let m_0 , \mathcal{A}_i , and k_j be as in the previous subsection. We write, for $n/2 \leq k \leq n$,

$$J_k = J_{k_0} + (J_{k_1} - J_{k_0}) + \cdots + (J_{k_N} - J_{k_{N-1}}). \quad (7.26)$$

If $\max_{n/2 \leq k \leq n} J_k \geq (1 + \varepsilon)\Delta^{-1}n \log \log \log n$, then either

- (a) $J_{k_0} \geq (1 + \frac{\varepsilon}{2})\Delta^{-1}n \log \log \log n$ for some $k_0 \in \mathcal{A}_0$, or else
- (b) for some $i \geq 1$ and k_i, k'_i consecutive elements of \mathcal{A}_i we have

$$J_{k'_i} - J_{k_i} \geq \frac{\varepsilon}{40i^2} \Delta^{-1}n \log \log \log n. \quad (7.27)$$

There are at most n/m_0 elements of \mathcal{A}_0 . Using Theorem 1.2, the probability of (a) is bounded by

$$c_1 \frac{n}{m_0} e^{-(1+\frac{\varepsilon}{4}) \log \log n}. \quad (7.28)$$

To estimate the probability in (b), suppose j and k are consecutive elements of \mathcal{A}_i . There are at most $2^{i+1}(n/m_0)$ such pairs. We have

$$\begin{aligned} J_k - J_j &= -\overline{B}([j+1, k]_{<}^2) - \overline{B}([1, j] \times [j+1, k]) \\ &\leq -\overline{B}([j+1, k]_{<}^2) + \mathbb{E} B([1, j] \times [j+1, k]) \\ &\leq -\overline{B}([j+1, k]_{<}^2) + c_2 \sqrt{j} \sqrt{k-j}, \end{aligned} \quad (7.29)$$

as in the previous subsection. Provided n is large enough, $c_2 \sqrt{j} \sqrt{k-j} = c_2 \sqrt{j} \sqrt{2^{-i}m_0}$ will be less than $\frac{\varepsilon}{80i^2} \Delta^{-1}n \log \log \log n$ for all i . So in order for $J_k - J_j$ to be larger than $\frac{\varepsilon}{40i^2} \Delta^{-1}n \log \log \log n$, we must have $-\overline{B}([j+1, k]_{<}^2)$ larger than $\frac{\varepsilon}{80i^2} \Delta^{-1}n \log \log \log n$. We use

Theorem 1.2 to bound this. Then multiplying by the number of pairs and summing over i , the probability is (b) is bounded by

$$\sum_{i=1}^{\infty} 2^{i+1} \frac{n}{m_0} e^{-\frac{\varepsilon}{80i^2} \frac{n}{2^{-i}m_0} \log \log n} \leq c_3 \frac{n}{m_0} e^{-c_4(n/m_0) \log \log n}. \quad (7.30)$$

We choose m_0 to be the largest possible power of 2 such that $c_4(n/m_0) > 2$.

Combining (7.28) and (7.30), we see that if we set $q > 1$ close to 1, $n_\ell = [q^\ell]$, and

$$E_\ell = \left\{ \max_{n_\ell/2 \leq k \leq n_\ell} J_k \geq (1 + \varepsilon) \Delta^{-1} n_\ell \log \log \log n_\ell \right\}, \quad (7.31)$$

then $\sum_\ell \mathbb{P}(E_\ell)$ is finite. So by Borel-Cantelli, the event E_ℓ happens for a last time, almost surely. Exactly as in the previous subsection, taking q close enough to 1 and using the fact that ε is arbitrary leads to the upper bound. \square

The proof of the lower bound is fairly similar to the previous subsection. Let $n_\ell = [e^{\ell^{1+\delta}}]$. Theorem 1.2 and Borel-Cantelli tell us that F_ℓ will happen infinitely often, where

$$F_\ell = \left\{ -\overline{B}([n_{\ell-1} + 1, n_\ell]_{<}^2) \geq (1 - \varepsilon) \Delta^{-1} n_\ell \log \log \log n_\ell \right\}. \quad (7.32)$$

We have

$$J_{n_\ell} \geq -\overline{B}([n_{\ell-1} + 1, n_\ell]_{<}^2) + J_{n_{\ell-1}} - A(1, n_{\ell-1}; n_{\ell-1}, n_\ell). \quad (7.33)$$

By the upper bound,

$$J_{n_{\ell-1}} = O(n_{\ell-1} \log \log \log n_{\ell-1}) = o(n_\ell \log \log \log n_\ell). \quad (7.34)$$

By Lemma 7.1,

$$\mathbb{P}(B([1, n_{\ell-1}] \times [n_{\ell-1} + 1, n_\ell]) \geq \varepsilon n_\ell \log \log \log n_\ell) \leq c_1 \exp\left(-c_2 \frac{\varepsilon n_\ell \log \log \log n_\ell}{\sqrt{n_{\ell-1}} \sqrt{n_\ell - n_{\ell-1}}}\right). \quad (7.35)$$

This is summable in ℓ , so

$$\limsup_{\ell \rightarrow \infty} \frac{B([1, n_{\ell-1}] \times [n_{\ell-1} + 1, n_\ell])}{n_\ell \log \log \log n_\ell} \leq \varepsilon \quad (7.36)$$

almost surely. This is true for every ε , so the limsup is 0. Combining this with (7.34) and substituting in (7.33) completes the proof. \square

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