Maximal Local Time of a d-dimensional Simple Random Walk on Subsets.

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Abstract: Strong theorems are given for the maximal local time on balls and subspaces for the *d*-dimensional simple symmetric random walk.

1. Introduction and main results

Consider a simple symmetric random walk $\{\mathbf{S}_n\}_{n=1}^{\infty}$ starting at the origin $\mathbf{0}$ on the d-dimensional integer lattice $\mathcal{Z}_d$, i.e. $\mathbf{S}_0 = \mathbf{0}$, $\mathbf{S}_n = \sum_{k=1}^{n} \mathbf{X}_k$, $n = 1, 2, \ldots$, where $\mathbf{X}_k$, $k = 1, 2, \ldots$ are i.i.d. random variables with distribution

$$\mathbf{P}(\mathbf{X}_1 = \mathbf{e}_i) = \mathbf{P}(\mathbf{X}_1 = -\mathbf{e}_i) = \frac{1}{2d}, \qquad i = 1, 2, ..., d$$

and $\{\mathbf{e}_1, \mathbf{e}_2, \ldots, \mathbf{e}_d\}$ is a system of orthogonal unit vectors in $\mathcal{Z}_d$. Define the local time of the walk by

(1.1)
$$\xi^{(d)}(\mathbf{x}, n) := \#\{k : \ 0 < k \le n, \ \mathbf{S}_k = \mathbf{x}\}.$$

where $\mathbf{x}$ is any lattice point of $\mathcal{Z}_d$. The maximal local time of the walk is defined as

(1.2)
$$\xi^{(d)}(n) := \max_{\mathbf{x} \in \mathcal{Z}_d} \xi^{(d)}(\mathbf{x}, n).$$

The properties of $\xi^d(n)$ were extensively studied in the cases $d = 1$, $d = 2$ and $d \ge 3$ separately. For $d = 1$ the interested reader should consult the monograph of P. Révész [8]. In this paper we are interested in investigating the maximum local time for $d \ge 2$ in a restricted sense, namely we want to investigate the maximum on certain subsets of the state space. It is easy to see that these maximums depend on both of the size and the shape of the selected subset. We will only investigate two types of subsets: balls centered at the origin and subspaces.

1.1 Two dimension.

In what follows we present the most important results on local time for $d = 2$ which are relevant to our investigation.

Theorem A (Erdős and Taylor [6])

$$\lim_{n \to \infty} \mathbf{P}(\xi^{(2)}(\mathbf{0}, n) < x \log n) = 1 - e^{-\pi x}.$$

Let $f(x)$ resp. $g(x)$ be a decreasing resp. increasing function for which $f(x) \log x \uparrow \infty$, $g(x)(\log x)^{-1} \downarrow 0$. Then

$$\pi^{-1} g(n) \log n \le \xi^{(2)}(\mathbf{0}, n)$$

finitely often with probability one if and only if

$$\int_1^{\infty} \frac{g(x)}{x \log x} e^{-g(x)} \, dx < \infty,$$

2

and

$$f(n) \log n \geq \xi^{(2)}(\mathbf{0}, n)$$

finitely often with probability one if and only if

$$\int_1^\infty \frac{f(x)}{x \log x} \, dx < \infty.$$

Theorem B (Erdős and Taylor [6])

$$\frac{1}{4\pi} \leq \liminf_{n \to \infty} \frac{\xi^{(2)}(n)}{(\log n)^2} \leq \limsup_{n \to \infty} \frac{\xi^{(2)}(n)}{(\log n)^2} \leq \frac{1}{\pi} \qquad a.s.$$

They also conjectured that the upper bound in the above theorem is the correct limit. This was confirmed recently by Dembo *et al.* [4]. In fact they proved the following more general result about aperiodic random walks. Recall that a random walk in $\mathcal{Z}_d$ is aperiodic (sometimes called adapted) if the steps are not supported on a proper subgroup of $\mathcal{Z}_d$. Note that this definition is equivalent to that in Spitzer [9], p. 20, and in this sense the simple symmetric random walk is aperiodic. In the sequel we use this definition of aperiodicity.

Theorem C ([4]) *Let $\mathbf{S}_n = \sum_{k=1}^n \mathbf{X}_k$ be an aperiodic random walk with i.i.d. increments $\mathbf{X}_k \in \mathcal{Z}_2$ that satisfy $\mathbf{E}\mathbf{X}_1 = 0$ and $\mathbf{E}|\mathbf{X}_1|^m < \infty$ for all $m < \infty$. Denote by $\Gamma = \mathbf{E}\mathbf{X}\mathbf{X}'$ the covariance matrix of the increments, and write $\pi_\Gamma := 2\pi(\det\Gamma)^{1/2}$. Define the local time and maximum local time as in the simple walk case. Let $M(n, \alpha)$ denote the number of points in the set $\{\mathbf{x} : \xi(\mathbf{x}, n) \geq \alpha(\log n)^2\}$. Then*

$$\lim_{n \to \infty} \frac{\xi^{(2)}(n)}{(\log n)^2} = \pi_\Gamma^{-1} \qquad a.s.$$

and for $\alpha \in (0, \pi_\Gamma^{-1}]$

$$\lim_{n \to \infty} \frac{\log M(n, \alpha)}{\log n} = 1 - \alpha\pi_\Gamma \qquad a.s.$$

Moreover any (random) sequence $\{\mathbf{x}_n\}$ such that $\xi^{(2)}(\mathbf{x}_n, n)/\xi^{(2)}(n) \to 1$, must satisfy

$$\lim_{n \to \infty} \frac{\log |\mathbf{x}_n|}{\log n} = \frac{1}{2} \qquad a.s.$$

For the simple symmetric random walk $\pi_\Gamma = \pi$.

3

As the above results show, the local time of every fixed point is roughly around $\log n$ but the maximal local time is around $(\log n)^2$. This phenomenon suggests that taking the maximum local time on appropriate subsets, one might get orders in between $\log n$ and $(\log n)^2$. The following result of Auer is in fact telling us that a set having maximal local time much greater than $\log n$ needs to be quite big.

Theorem D (Auer [1]) *For any $\varepsilon > 0$ we have*

$$\lim_{n \to \infty} \sup_{||\mathbf{x}|| \le r_n} \left| \frac{\xi^{(2)}(\mathbf{x}, n)}{\xi^{(2)}(\mathbf{0}, n)} - 1 \right| = 0 \qquad a.s.,$$

where

$$r_n = \exp((\log n)^{1/2}(\log \log n)^{-1/2-\varepsilon})$$

and $||\mathbf{x}||$ stands for the usual Euclidean norm.

Now let A be a subset of $\mathcal{Z}_2$ and define

$$(1.3) \qquad\qquad \xi_A^{(2)}(n) := \max_{\mathbf{x} \in A} \xi^{(2)}(\mathbf{x}, n).$$

Let moreover $B(r)$ denote the set of lattice points in the disc of radius r centered at the origin, i.e.
$$(1.4) \qquad\qquad B(r) := \{\mathbf{x} \in \mathcal{Z}_2 : ||\mathbf{x}|| \le r\}.$$

Then with r_n as in Theorem D, for any subset $A \subseteq B(r_n)$, $\xi_A^{(2)}(n)/\xi^{(2)}(\mathbf{0}, n) \to 1$ as $n \to \infty$. Hence for $\xi_A^{(2)}(n)$ we have the same asymptotic results as for $\xi^{(2)}(\mathbf{0}, n)$.

Denote by $L = L(a_1, a_2)$ the lattice points $\mathbf{x} = (x_1, x_2)$ on the line $a_1 x_1 + a_2 x_2 = 0$, where a_1 and a_2 are integers not both of them zero. Now we formulate our results for the two-dimensional case. Our first theorem is telling us that for any line going through the origin which contains lattice points at all, the maximal local time has the same order of magnitude as for the whole plane.

Theorem 1.1. *For any line $L = L(a_1, a_2)$ such that a_1, a_2 are integers, not both of them zero, we have*
$$(1.5) \qquad\qquad \frac{1}{8\pi} \le \liminf_{n \to \infty} \frac{\xi_L^{(2)}(n)}{(\log n)^2} \le \limsup_{n \to \infty} \frac{\xi_L^{(2)}(n)}{(\log n)^2} \le \frac{1}{2\pi} \qquad a.s.$$

The next two theorems contain our results about discs centered at the origin.

4

Theorem 1.2. *Let* $r_n = n^\alpha$, $0 < \alpha \le 1/2$. *Then*

$$(1.6) \qquad \frac{4\alpha^2}{\pi} \le \liminf_{n\to\infty} \frac{\xi^{(2)}_{B(r_n)}(n)}{(\log n)^2} \le \limsup_{n\to\infty} \frac{\xi^{(2)}_{B(r_n)}(n)}{(\log n)^2} \le \frac{2\alpha}{\pi} \qquad a.s.$$

Remark. The upper part of (1.6) is not new. It was proved by Dembo *et al.* [4] as a byproduct of their proof of Theorem 5.1 (see page 17 in their paper).

Theorem 1.3. *Let* $r_n = \exp((\log n)^\beta)$, *where* $1/2 \le \beta < 1$. *Then*

$$(1.7) \qquad \frac{4(1-\varepsilon)}{\pi}(\log n)^{2\beta} \le \xi^{(2)}_{B(r_n)}(n) \le (\log n)^{2\beta+\varepsilon} \qquad a.s.$$

for any $\varepsilon > 0$ *and for all but finitely many* n.

Corollary 1.1 *If* L *and* $B(r_n)$ *are the sets in Theorems* 1.1, 1.2 *and* 1.3 *respectively, then for* $r_n = n^\alpha$, $0 < \alpha \le 1/2$ *we have*

$$(1.8) \qquad \frac{\alpha^2}{2\pi} \le \liminf_{n\to\infty} \frac{\xi^{(2)}_{B(r_n)\cap L}(n)}{(\log n)^2} \le \limsup_{n\to\infty} \frac{\xi^{(2)}_{B(r_n)\cap L}(n)}{(\log n)^2} \le \frac{1}{\pi}\min\left(\frac{1}{2}, 2\alpha\right) \qquad a.s.$$

and for $r_n = \exp((\log n)^\beta)$ *with* $1/2 \le \beta < 1$ *we have*

$$(1.9) \qquad \frac{1-\varepsilon}{2\pi}(\log n)^{2\beta} \le \xi^{(2)}_{B(r_n)\cap L}(n) \le (\log n)^{2\beta+\varepsilon} \qquad a.s.$$

for any $\varepsilon > 0$ *and for all but finitely many* n.

1.2 Three and higher dimension.

Just like in two dimension, for a subset $A \subseteq \mathcal{Z}_d$ we define

$$(1.10) \qquad \xi^{(d)}_A(n) := \max_{\mathbf{x}\in A} \xi^{(d)}(\mathbf{x}, n).$$

To formulate the most important known results on $\xi^{(d)}(n)$ of (1.2), we need some more definition. Denote by $\gamma_d(n)$ the probability that in the first $n-1$ steps the path does not return to the origin. Then

$$1 = \gamma_d(1) \ge \gamma_d(2) \ge ... \ge \gamma_d(n) \ge ... > 0.$$

It was proved in Dvoretzky and Erdős [5] that, for $d \geq 3$

$$(1.11) \qquad \gamma_d(n) \to \gamma_d > 0,$$

and

$$(1.12) \qquad \gamma_d < \gamma_d(n) < \gamma_d + O(n^{1-d/2})$$

as $n \to \infty$. So γ_d is the probability that the d-dimensional simple symmetric random walk never returns to its starting point.

Let $\xi^{(d)}(\mathbf{0}, \infty)$ be the total local time at 0 of the infinite path in $\mathcal{Z}_d$. Then (see [6]) $\xi^{(d)}(\mathbf{0}, \infty)$ has geometric distribution:

$$(1.13) \qquad \mathbf{P}(\xi^{(d)}(\mathbf{0}, \infty) = k) = \gamma_d(1 - \gamma_d)^k, \qquad k = 0, 1, 2, \ldots$$

Erdős and Taylor [6] proved the following strong law for the maximal local time:

Theorem F ([6]) *For $d \geq 3$*

$$(1.14) \qquad \lim_{n \to \infty} \frac{\xi^{(d)}(n)}{\log n} = \lambda_d \qquad a.s.,$$

where

$$(1.15) \qquad \lambda_d = -\frac{1}{\log(1 - \gamma_d)}.$$

Remark. For the exact value of γ_3 see e.g. Spitzer [9], p. 103 which implies that $\lambda_3 < 1$ and hence $\lambda_d < 1$ for all $d \geq 3$.

Let $B(r)$ stand for the (discrete) ball centered at the origin in the d-dimensional space and having radius r, i.e.

$$(1.16) \qquad B(r) := \{\mathbf{x} \in \mathcal{Z}_d : \|\mathbf{x}\| \leq r\}.$$

Let furthermore $\mathbf{x} = (x_1, x_2, \ldots, x_d)$,

$$S_{d-1} := \{\mathbf{x} \in \mathcal{Z}_d : a_1 x_1 + a_2 x_2 + \ldots + a_d x_d = 0\}$$

and

$$S_{d-2} := \{\mathbf{x} \in \mathcal{Z}_d : a_1 x_1 + a_2 x_2 + \ldots + a_d x_d = 0, \quad b_1 x_1 + b_2 x_2 + \ldots + b_d x_d = 0\}$$

with integer coefficients $a_1, a_2, \ldots a_d, b_1, b_2, \ldots b_d$.

For subspaces we will prove the following two results.

6

Theorem 1.4. *Let $d \geq 3$. Suppose that $a_1, a_2, ...a_d$ are integers, not all of them zero, then*

$$\lim_{n\to\infty} \frac{\xi_{S_{d-1}}^{(d)}(n)}{\log n} = \frac{\lambda_d}{2} \qquad a.s.$$

The above theorem is telling us that the maximal local time in the $d-1$ dimensional subspace has the same order of magnitude as in the whole d-dimensional space. On the other hand, the next theorem shows that in the $d-2$ dimensional subspace the maximal local time gets drastically smaller. Note that in $d-3$-dimensional subspace the maximal local time remains finite with probability 1 as $n \to \infty$.

Theorem 1.5. *Let $d \geq 3$. Suppose that $a_1, a_2, ...a_d$ are integers, not all of them zero and $b_1, b_2, ...b_d$ are also integers not all of them zero. Assume also that the vectors $(a_1, a_2, \ldots, a_d)$ and $(b_1, b_2, \ldots, b_d)$ are not parallel. Then*

$$\lim_{n\to\infty} \frac{\xi_{S_{d-2}}^{(d)}(n)}{\log \log n} = \lambda_d \qquad a.s.$$

For balls centered at the origin we will prove the following result:

Theorem 1.6. *Let $d \geq 3$. For any sequence $r_n \uparrow \infty$, such that $\limsup_{n\to\infty}(\log r_n)/(\log n) \leq 1/2$, we have*

(1.17) $$\lim_{n\to\infty} \frac{\xi_{B(r_n)}^{(d)}(n)}{\log r_n} = 2\lambda_d \qquad a.s.$$

The organization of the paper is as follows. In Section 2 we will present some well-known facts and prove some preliminary results. Sections 3 and 4 contain the proofs of the two dimensional and higher dimensional results, respectively. In Section 5 some implications of the above results and some open questions will be discussed. Throughout the paper $c, c_1, \ldots, C, C_1, \ldots$ will denote positive constants, the value of which is unimportant and may vary from line to line.

2. Preliminary facts and results

First we present the d-dimensional law of the iterated logarithm and rate of escape for simple walk.

Fact 1. (Dvoretzky and Erdős [5] or [8], pp. 193, 195) *For a simple symmetric random walk in $\mathcal{Z}_d$*

$$(2.1) \qquad \limsup_{n\to\infty}(2n\log\log n)^{-1/2}d^{1/2}||\mathbf{S}_n|| = 1 \qquad a.s.$$

Moreover, in case $d \geq 3$, for any nonincreasing nonnegative sequence $b(n)$ we have

$$(2.2) \qquad n^{1/2}b(n) \geq d^{1/2}||\mathbf{S}_n|| \qquad a.s.$$

for all but finitely many n, if and only if

$$(2.3) \qquad \sum_{n=1}^{\infty}(b(2^n))^{d-2} < \infty.$$

Consider now the case $d = 2$. Introduce

$$
\begin{aligned}
(2.4) \qquad p(\mathbf{x}) : \ &= \ \mathbf{P}(\min\{n > 0 : \mathbf{S}_n = \mathbf{0}\} > \min\{n > 0 : \mathbf{S}_n = \mathbf{x}\}) \\
&= \ \mathbf{P}(\{\mathbf{S}_n\} \text{ reaches } \mathbf{x} \text{ before returning to } \mathbf{0}).
\end{aligned}
$$

We will need the following two lemmas from Révész [8]:

Fact 2. ([8], p. 219) *For a simple symmetric random walk in $\mathcal{Z}_2$ there exists a positive constant C such that for any $\mathbf{x} \in \mathcal{Z}_2$ with $||\mathbf{x}|| \geq 2$*

$$(2.5) \qquad p(\mathbf{x}) \geq \frac{C}{\log ||\mathbf{x}||}.$$

Let us define

$$(2.6) \qquad \rho_0 := 0, \qquad \rho_n := \min\{k : k > \rho_{n-1}, \mathbf{S}_k = \mathbf{0}\}, \quad n = 1, 2, \ldots$$

Fact 3. ([8], pp. 219-220.) *For a simple symmetric random walk in $\mathcal{Z}_2$ let*

$$(2.7) \qquad Y_i(\mathbf{x}) := \xi^{(2)}(\mathbf{x}, \rho_i) - \xi^{(2)}(\mathbf{x}, \rho_{i-1}), \quad i = 1, 2, \ldots$$

Then for fixed $\mathbf{x} \in \mathcal{Z}_2$, $\{Y_i(\mathbf{x})\}_{i=1}^{\infty}$ are i.i.d. random variables with the following distribution:

$$(2.8) \qquad \mathbf{P}(Y_1(\mathbf{x}) = 0) = 1 - p(\mathbf{x})$$

$$(2.9) \qquad \mathbf{P}(Y_1(\mathbf{x}) = k) = (1 - p(\mathbf{x}))^{k-1}p^2(\mathbf{x}), \quad k = 1, 2, \ldots$$

Now we will prove our first

8

Lemma 2.1. *For a simple symmetric random walk in $\mathcal{Z}_2$ we have for arbitrary $\mathbf{x} \in \mathcal{Z}_2$, and any $u > 2$*

$$(2.10) \qquad \mathbf{P}\left(\sum_{k=1}^{n} Y_k(\mathbf{x}) > u\,n\right) \leq e^{np(\mathbf{x})(1-u/2)}.$$

Remark. Fact 3 and Lemma 2.1 are true for more general random walk, but we need them only for simple symmetric case.

Proof: From (2.9) we easily get with $q(\mathbf{x}) = 1 - p(\mathbf{x})$ that for any $z > 0$, for which $q(\mathbf{x})e^z < 1$,

$$(2.11) \qquad \mathbf{E}\left(e^{zY_1(\mathbf{x})}\right) = q(\mathbf{x}) + \frac{p^2(\mathbf{x})e^z}{1 - q(\mathbf{x})e^z}.$$

Putting $z = \log(2/(1 + q(\mathbf{x})))$, we have

$$(2.12) \qquad q(\mathbf{x})e^z = \frac{2q(\mathbf{x})}{1 + q(\mathbf{x})} < 1$$

so (2.11) holds. Thus

$$(2.13) \qquad \mathbf{E}\left(e^{zY_1(\mathbf{x})}\right) = q(\mathbf{x}) + \frac{p^2(\mathbf{x})e^z}{1 - q(\mathbf{x})e^z} = 1 + p(\mathbf{x}).$$

By exponential Markov inequality, and (2.13) we have

$$\mathbf{P}\left(\sum_{k=1}^{n} Y_k(\mathbf{x}) > u\,n\right) \quad \leq \quad \frac{\left(\mathbf{E}e^{zY_1(\mathbf{x})}\right)^n}{e^{zun}} = (1 + p(\mathbf{x}))^n \left(1 - \frac{p(\mathbf{x})}{2}\right)^{un}$$

$$(2.14) \qquad\qquad\qquad\qquad \leq \quad \exp\left(np(\mathbf{x})(1 - u/2)\right),$$

where the inequality $1 + v \leq e^v$ was used. $\square$

For our next lemma we need further notations and facts. Our main source for these is Spitzer's book [9]. Here we consider a two-dimensional symmetric aperiodic recurrent walk on $\mathcal{Z}_2$, more general than a simple symmetric random walk which will be denoted by $\{\mathbf{S}_n\}$ as well. All what we are quoting however work in case of a one-dimensional walk under the same constraints as well. We adopt the notations and definitions listed below from [9]. Denote the n-step probability transition function by

$$P_n(\mathbf{x}, \mathbf{y}) = P_n(\mathbf{y}, \mathbf{x}) = P_n(\mathbf{0}, \mathbf{x} - \mathbf{y}) = \mathbf{P}(\mathbf{S}_n = \mathbf{x} - \mathbf{y}|\mathbf{S}_0 = \mathbf{0}), \quad n = 0, 1, \ldots, \quad \mathbf{x}, \mathbf{y} \in \mathcal{Z}_2.$$

9

For $b \geq 0$ integer we define

$$(2.15) \qquad\qquad \mathbf{T}_b := \min\{j > 0 : \mathbf{S}_{b+j} = \mathbf{0}\}$$

and will denote $\mathbf{T}_0 =: \mathbf{T}$. Let (see [9], pp. 107, 160-161)

$$(2.16) \qquad \mathbf{Q}_n(\mathbf{x}, \mathbf{y}) := \mathbf{P}_{\mathbf{x}}(\mathbf{S}_n = \mathbf{y}, \ \mathbf{T} > n), \quad n = 0, 1, \dots, \quad \mathbf{x}, \mathbf{y} \in \mathcal{Z}_2 - \{\mathbf{0}\},$$

and

$$(2.17) \qquad\qquad g(\mathbf{x}, \mathbf{y}) := \sum_{n=0}^{\infty} \mathbf{Q}_n(\mathbf{x}, \mathbf{y}), \quad \mathbf{x}, \mathbf{y} \in \mathcal{Z}_2 - \{\mathbf{0}\}$$

where $\mathbf{P}_{\mathbf{x}}(\cdot) := \mathbf{P}(\cdot | \mathbf{S}_0 = \mathbf{x})$. Using $\mathbf{P}(\cdot) := \mathbf{P}_{\mathbf{0}}(\cdot)$ we also recall from Spitzer [9] that

$$(2.18) \quad \mathbf{P}_{\mathbf{x}}(\mathbf{T} > n) = \sum_{\mathbf{t} \neq \mathbf{0}} g(\mathbf{x}, \mathbf{t}) \sum_{\mathbf{y} \neq \mathbf{0}} \mathbf{Q}_n(\mathbf{t}, \mathbf{y}) P_1(\mathbf{y}, \mathbf{0}), \quad n = 0, 1, \dots, \ \mathbf{x} \in \mathcal{Z}_2 - \{\mathbf{0}\}$$

and introducing

$$(2.19) \quad v_n(\mathbf{t}) = \frac{1}{\mathbf{P}(\mathbf{T} > n + 1)} \sum_{\mathbf{y} \neq \mathbf{0}} \mathbf{Q}_n(\mathbf{t}, \mathbf{y}) P_1(\mathbf{y}, \mathbf{0}), \quad n = 0, 1, \dots, \ \mathbf{t} \in \mathcal{Z}_2 - \{\mathbf{0}\},$$

we get for $\mathbf{x} \neq \mathbf{0}$

$$(2.20) \qquad \mathbf{P}_{\mathbf{x}}(\mathbf{T} > n) = \mathbf{P}(\mathbf{T} > n + 1) \sum_{\mathbf{t} \neq \mathbf{0}} g(\mathbf{x}, \mathbf{t}) v_n(\mathbf{t}), \quad n = 0, 1, \dots$$

with

$$(2.21) \qquad\qquad v_n(\mathbf{t}) \geq 0, \qquad \sum_{\mathbf{t} \neq \mathbf{0}} v_n(\mathbf{t}) = 1.$$

We introduce further the notations (see [9], pp. 114, 139, 328)

$$(2.22) \qquad\qquad \mathbf{G}_n(\mathbf{x}, \mathbf{y}) = \sum_{k=0}^{n} P_k(\mathbf{x}, \mathbf{y}), \quad n = 0, 1, \dots, \ \mathbf{x}, \mathbf{y} \in \mathcal{Z}_2,$$

the truncated Green function

$$(2.23) \qquad\qquad h(n) = \mathbf{G}_n(\mathbf{0}, \mathbf{0}) = \sum_{k=0}^{n} P_k(\mathbf{0}, \mathbf{0}), \quad n = 0, 1, \dots$$

and the potential kernel

$$(2.24) \qquad\qquad a(\mathbf{x}) = \sum_{n=0}^{\infty} (P_n(\mathbf{0}, \mathbf{0}) - P_n(\mathbf{0}, \mathbf{x})), \quad \mathbf{x} \in \mathcal{Z}_2.$$

Then it is known (see Spitzer [9], p. 355) that

$$(2.25) \qquad 0 \le g(\mathbf{x}, \mathbf{y}) = a(\mathbf{x}) + a(\mathbf{y}) - a(\mathbf{x} - \mathbf{y}).$$

We recall from Spitzer [9], p. 139

Fact 4. *For any symmetric recurrent aperiodic random walk in two dimension*

$$(2.26) \qquad \sum_{\mathbf{y} \in \mathcal{Z}_2} P_{n+1}(\mathbf{x}, \mathbf{y}) a(\mathbf{y}) = a(\mathbf{x}) + \mathbf{G}_n(\mathbf{x}, \mathbf{0}), \quad n = 0, 1, \dots, \mathbf{x} \in \mathcal{Z}_2.$$

Now we are ready to prove our

Lemma 2.2. *For any symmetric, recurrent aperiodic walk in two dimension we have*

$$(2.27) \qquad \mathbf{P}(\mathbf{T}_b \ge k) \le 2 \frac{h(b)}{h(k-1)}, \quad b = 0, 1, \dots, k = 2, 3, \dots$$

Proof: We start with the following simple observation: for $k \ge 1$, $b \ge 0$

$$(2.28) \qquad \mathbf{P}(\mathbf{T}_b \ge k) = \sum_{\mathbf{x} \in \mathcal{Z}_2} \mathbf{P}(\mathbf{T}_b \ge k | \mathbf{S}_b = \mathbf{x}) \mathbf{P}(\mathbf{S}_b = \mathbf{x}) = \sum_{\mathbf{x} \in \mathcal{Z}_2} P_b(\mathbf{0}, \mathbf{x}) \mathbf{P}_{\mathbf{x}}(\mathbf{T} \ge k).$$

Now (2.25) and the symmetry of the walk implies that

$$(2.29) \qquad a(\mathbf{x} + \mathbf{y}) \le a(\mathbf{x}) + a(\mathbf{y}).$$

Moreover, denoting $\mathbf{x} + \mathbf{y} = \mathbf{z}$ we easily get from (2.29) that

$$(2.30) \qquad a(\mathbf{z}) \le a(\mathbf{x}) + a(\mathbf{z} - \mathbf{x}) = a(\mathbf{x}) + a(\mathbf{x} - \mathbf{z}),$$

thus
$$(2.31) \qquad a(\mathbf{z}) - a(\mathbf{x} - \mathbf{z}) \le a(\mathbf{x}),$$

which in turn, combined with (2.25) implies that

$$(2.32) \qquad g(\mathbf{x}, \mathbf{y}) \le 2\, a(\mathbf{x}).$$

Combining (2.20), (2.21) and (2.32) we conclude that for $\mathbf{x} \ne \mathbf{0}$

$$(2.33) \qquad \mathbf{P}_{\mathbf{x}}(\mathbf{T} > n) \le 2a(\mathbf{x}) \mathbf{P}(\mathbf{T} > n + 1), \quad n = 0, 1, \dots$$

11

Now using (2.28), (2.33), Fact 4 and the simple observation that $a(\mathbf{0}) = 0$, we get that

$$\mathbf{P}(\mathbf{T}_b \geq k) \leq 2\mathbf{P}(\mathbf{T} \geq k+1)\sum_{\mathbf{x}\in\mathcal{Z}_2-\{\mathbf{0}\}} P_b(\mathbf{0},\mathbf{x})a(\mathbf{x}) + P_b(\mathbf{0},\mathbf{0})\mathbf{P}(\mathbf{T}\geq k)$$

$$\leq 2\mathbf{P}(\mathbf{T}\geq k)\left(\sum_{\mathbf{x}\in\mathcal{Z}_2} P_b(\mathbf{0},\mathbf{x})a(\mathbf{x}) + P_b(\mathbf{0},\mathbf{0})\right)$$

$$(2.34) \qquad \leq 2\mathbf{P}(\mathbf{T}\geq k)(\mathbf{G}_{b-1}(\mathbf{0},\mathbf{0}) + P_b(\mathbf{0},\mathbf{0})) \leq 2\mathbf{P}(\mathbf{T}\geq k)h(b).$$

To estimate $\mathbf{P}(\mathbf{T} > k)$ we use an argument essentially from Erdős and Taylor [6]. Partitioning according to the last return to zero we get

$$(2.35) \qquad \sum_{j=0}^{k} \mathbf{P}(\mathbf{T} > k-j)P_j(\mathbf{0},\mathbf{0}) = 1$$

implying that

$$(2.36) \qquad \mathbf{P}(\mathbf{T} > k)\sum_{j=0}^{k} P_j(\mathbf{0},\mathbf{0}) \leq 1,$$

hence
$$(2.37) \qquad \mathbf{P}(\mathbf{T} > k) \leq 1/h(k).$$

Now (2.34) and (2.37) imply our lemma. $\square$

Fact 5. *If the walk is recurrent and aperiodic with finite second moment, then we have as* $n \to \infty$

$$(2.38) \qquad h(n) \sim c_2 \log n \qquad \text{for } d = 2,$$

$$(2.39) \qquad h(n) \sim c_1\sqrt{n} \qquad \text{for } d = 1.$$

The case $d = 2$ is well-known for strongly aperiodic walk (see [9], p. 75 and [4] how to weaken this condition for the aperiodic case). The case $d = 1$ is well-known (see e.g. [9], p. 381).

Let $\{\mathbf{S}_n\}$ be a symmetric recurrent aperiodic walk in two dimension. Consider the following problem. Let $\mathbf{T}_b$ be defined by (2.15). At $b + \mathbf{T}_b$ the walk is at $\mathbf{0}$. Now after another b steps we wait again until the walk returns to $\mathbf{0}$. Keep repeating this procedure, we would like to estimate the number of such returns within n steps. This problem was considered in [2] in one dimension. Here we repeat essentially the same argument and spell it out in the

12

$d = 2$ case with the appropriate modifications, using Lemma 2.2. To formulate this problem precisely, let

$$(2.40) \quad \zeta_1(b) := b, \quad \alpha_k(b) := \mathbf{T}_{\zeta_k(b)}, \quad \zeta_{k+1}(b) := \sum_{i=1}^{k} \alpha_i(b) + (k+1)b, \quad k = 1, 2, \ldots$$

Then $\alpha_k(b)$, $k = 1, 2, \ldots$ are i.i.d. random variables having the same distribution as $\mathbf{T}_b$.

Lemma 2.3. *For a symmetric recurrent aperiodic random walk in $\mathcal{Z}_2$ we have for $b > 1$, $u > 1$, $k \geq 1$*

$$(2.41) \qquad\qquad \mathbf{P}\left(\sum_{i=1}^{k} \alpha_i(b) \geq u \right) \leq C\, k \frac{\log b}{\log u}.$$

Proof: By Lemma 2.2 and Fact 5

$$(2.42) \qquad\qquad \mathbf{P}(\mathbf{T}_b \geq k) \leq 2\frac{h(b)}{h(k-1)} \leq C\frac{\log b}{\log(k-1)} \leq C\frac{\log b}{\log k}.$$

Let

$$(2.43) \qquad\qquad {}^u\alpha_k(b) = \begin{cases} \alpha_k(b) & \text{if} \quad \alpha_k(b) \leq u \\ 0 & \text{if} \quad \alpha_k(b) > u \end{cases}$$

First observe, that

$$(2.44) \qquad \mathbf{E}({}^u\alpha_i(b)) \leq \sum_{j=0}^{u} \mathbf{P}(\mathbf{T}_b \geq j) \leq 2 + C\sum_{j=2}^{u} \frac{\log b}{\log j} \leq Cu\frac{\log b}{\log u}.$$

Define the event

$$(2.45) \qquad\qquad A = \bigcap_{j=1}^{k} \{\alpha_j(b) < u\}.$$

$$
\begin{aligned}
(2.46) \qquad \mathbf{P}\left(\sum_{i=1}^{k} \alpha_i(b) \geq u \right) \;\leq\; & \mathbf{P}\left(\sum_{i=1}^{k} \alpha_i(b) \geq u, A \right) + \mathbf{P}(\bar{A}) \leq \mathbf{P}\left(\sum_{i=1}^{k} {}^u\alpha_i(b) \geq u \right) \\
& + k\mathbf{P}(\alpha_i(b) \geq u) \leq \frac{k\mathbf{E}({}^u\alpha_i(b))}{u} + kC\frac{\log b}{\log u} \leq Ck\frac{\log b}{\log u}.
\end{aligned}
$$

$\square$

Let $b_t > 1$ be an integer valued function of t and let ν_t be the largest integer N for which

$$(2.47) \qquad\qquad \sum_{i=1}^{N} \alpha_i(b_t) + (N+1)b_t \leq t.$$

13

Lemma 2.4. *Under the conditions of Lemma 2.3 for $f(t) > 0$, $(f(t) + 2)b_t < t - 1$ we have*

$$(2.48) \qquad \mathbf{P}(\nu_t \le f(t)) \le Cf(t) \frac{\log b_t}{\log(t - (f(t) + 2)b_t)}.$$

Proof: By Lemma 2.3

$$\mathbf{P}(\nu_t \le f(t)) \quad \le \quad \mathbf{P}(\nu_t \le [f(t)] + 1) \le \mathbf{P}\left(\sum_{i=1}^{[f(t)]+1} \alpha_i(b_t) > t - ([f(t)] + 2)b_t) \right)$$

$$(2.49) \qquad \qquad \le \quad C(f(t) + 1) \frac{\log b_t}{\log(t - (f(t) + 2)b_t)},$$

implying our statement. □

As mentioned above, the quantities defined in this section have analogues in one dimension. Also, there are corresponding one-dimensional analogues of results quoted for the two-dimensional case. We recall the one-dimensional versions of Lemmas 2.3 and 2.4.

Fact 6. ([2]) *Consider a symmetric aperiodic random walk $\{S_n\}$ on $\mathcal{Z}_1$ with finite variance and define $\alpha_i(b)$ and ν_t exactly as before for $\{S_n\}$. Then for $b > 1$, $u > 1$, $k \ge 1$*

$$(2.50) \qquad \mathbf{P}\left(\sum_{i=1}^{k} \alpha_i(b) \ge u \right) \le C k \sqrt{\frac{b}{u}}.$$

Furthermore, if $f(t) > 0$, $b_t > 0$, $(f(t) + 2)b_t < t$ we have

$$(2.51) \qquad \mathbf{P}(\nu_t \le f(t)) \le Cf(t) \frac{\sqrt{b_t}}{\sqrt{t - (f(t) + 2)b_t}}.$$

We will need the following upper tail estimates essentially from Erdős and Taylor [6].

Fact 7. ([6] or [4]) *For the simple symmetric random walk on the plane we have for any $\alpha > 0$ and $0 < \delta < 1$, $n \ge 1$*

$$(2.52) \qquad \mathbf{P}(\xi^{(2)}(\mathbf{0}, n) \ge \alpha(\log n)^2) < n^{-(1-\delta)\pi\alpha}.$$

Fact 8. ([6], (3.6)) *For the simple symmetric random walk on the plane we have for any $\alpha > 0$, $\delta > 0$, $n \ge 1$*

$$(2.53) \qquad \mathbf{P}(\xi^{(2)}(\mathbf{0}, n) \ge \alpha(\log n)^2) > n^{-(1+\delta)\pi\alpha}.$$

14

Fact 9. ([2]) *Let $\{S_n\}$ be a one-dimensional symmetric aperiodic random walk on $\mathcal{Z}_1$ with $\sigma^2 = E(X_1^2) < \infty$ and suppose that $x_n \to \infty$, $x_n/n^{1/2} \to 0$ as $n \to \infty$. Let $\xi(0,n)$ be its local time at zero. Then for any $\varepsilon > 0$ and large enough n*

$$(2.54) \quad C_1 \exp\left(-(1+\varepsilon)\frac{x_n^2\sigma^2}{2}\right) \leq \mathbf{P}(\xi(0,n) \geq x_n \, n^{1/2}) \leq C_2 \exp\left(-(1-\varepsilon)\frac{x_n^2\sigma^2}{2}\right).$$

Fact 10. ([3], (2.1) and [7], Lemma 2.5.) *For a symmetric aperiodic random walk with finite variance in $\mathcal{Z}_2$ we have for any $x > 0$*

$$(2.55) \qquad\qquad \mathbf{P}\left(\xi^{(2)}(\mathbf{0},n) \geq x \log n\right) \leq \exp(-cx)$$

with some constant $c > 0$.

3. Proofs of the two dimensional results.

Now let $\{\mathbf{S}_n\}$ be a two-dimensional simple symmetric random walk as defined in the Introduction with steps $\{\mathbf{X}_i\}$.

Proof of Theorem 1.1. Consider the line $L = L(a_1, a_2)$ with $a_1 x_1 + a_2 x_2 = 0$, where a_1 and a_2 are integers, not both of them zero. Without loss of generality we may assume that a_1 and a_2 are relatively prime.

For the two-dimensional random walk define a one-dimensional walk with steps

$$Y_i = a_j \quad \text{if} \quad \mathbf{X}_i = \mathbf{e}_j, \qquad i = 1, 2, \dots \quad j = 1, 2.$$

$$Y_i = -a_j \quad \text{if} \quad \mathbf{X}_i = -\mathbf{e}_j, \qquad i = 1, 2, \dots \quad j = 1, 2.$$

Then $Z_n = \sum_{i=1}^n Y_i$ is an aperiodic one-dimensional symmetric random walk with $Z_n = 0$ if and only if $\mathbf{S}_n \in L(a_1, a_2)$. Thus denoting by $V^L(n) = \#\{i : 1 \leq i \leq n, \ \mathbf{S}_i \in L(a_1, a_2)\}$, the number of visits of $\mathbf{S}_n$ on $L(a_1, a_2)$, we have

$$(3.1) \qquad\qquad V^L(n) = \xi^Z(0,n),$$

where $\xi^Z(\cdot, n)$ is the local time of $\{Z_n\}$.

To get the upper bound in our theorem, denote by $D(n)$ the set of lattice points on L which are visited up to n by $\{\mathbf{S}_i\}$, and denote by $|D(n)|$ the number of points in $D(n)$. Select a subsequence $n_j = [e^j]$, $j = 1, 2, \dots$ and let for any $\mathbf{x} \in \mathcal{Z}_2$

$$C_{\mathbf{x}}^j = \{\xi^{(2)}(\mathbf{x}, n_{j+1}) > \lambda(\log n_j)^2\}.$$

15

Then using (3.1), Facts 7 and 9 we conclude that for any $\lambda > 0$, $\delta > 0$, $\varepsilon > 0$ and j large enough

$$\mathbf{P}(\xi_L^{(2)}(n_{j+1}) > \lambda(\log n_j)^2)$$

$$\leq \ \mathbf{P}\left(\bigcup_{\mathbf{x} \in D(n_{j+1})} C_{\mathbf{x}}^j, \ |D(n_{j+1})| \leq \sqrt{n_{j+1}\log n_{j+1}}\right) + \mathbf{P}\left(|D(n_{j+1})| > \sqrt{n_{j+1}\log n_{j+1}}\right)$$

$$\leq \ \mathbf{P}\left(\bigcup_{\mathbf{x} \in D(n_{j+1})} C_{\mathbf{x}}^j, \ |D(n_{j+1})| \leq \sqrt{n_{j+1}\log n_{j+1}}\right) + \mathbf{P}\left(\xi^Z(0,n_{j+1}) > \sqrt{n_{j+1}\log n_{j+1}}\right)$$

$$\leq \ \sqrt{n_{j+1}\log n_{j+1}} \ \mathbf{P}(\xi^{(2)}(\mathbf{0},n_{j+1}) > \lambda(\log n_j)^2) + C_2\exp\left(-(j+1)\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

$$\leq \ \sqrt{n_{j+1}\log n_{j+1}} \ \exp\left(-\lambda(1-\delta)\pi\frac{j^2}{j+1}\right) + C_2\exp\left(-j\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

$$(3.2) \quad \leq \exp\left(-(1-\delta)^2\pi\,j\,\lambda + j(1/2+\varepsilon)\right) + C_2\exp\left(-j\frac{(1-\varepsilon)\sigma^2}{2}\right),$$

where we used in the second inequality above that $|D(n)| \leq V^L(n)$. Here σ^2 is the variance of Y_i and hence depends only on a_1 and a_2. It is easy to see that we can choose $\varepsilon > 0$, $\delta > 0$ for which the last line of (3.2) is summable in j whenever $\lambda > 1/(2\pi)$, which in turn, using Borel-Cantelli lemma and the usual monotonicity argument implies

$$(3.3) \qquad\qquad \limsup_{n\to\infty} \frac{\xi_L^{(2)}(n)}{(\log n)^2} \leq \frac{1}{2\pi} \qquad a.s.,$$

so we have the upper half of the theorem.

To get the lower bound, we will essentially follow Erdős and Taylor's argument with some modification. We consider the walk $\{\mathbf{S}_i\}$, wait $[n^\alpha]$ steps and observe $\xi^{(2)}(\mathbf{0}, [n^\alpha])$. The number $0 < \alpha < 1$ will be chosen later. At time $[n^\alpha]$ the walk is somewhere on the plane, and we wait until its first return to $L = L(a_1, a_2)$. When it returns to L, we observe the local time of the hitting point of L for a time interval $[n^\alpha]$, and then wait again for the walk to return to L. We keep repeating this procedure for a total time of n steps. This construction ensures that the local times of these hitting points over a time interval $[n^\alpha]$ are i.i.d. random variables having the same distribution as $\xi^{(2)}(\mathbf{0}, [n^\alpha])$. Combining this observation with our Fact 6, would produce our lower bound in the theorem. In what follows we work out the above outlined ideas with the added complexity of working with subsequences as before. Let

16

$n_j = j^\beta$, $j = 1, 2, \ldots$ with integer β to be chosen later. Now fix $j \geq 1$ and define for $k = 1, 2, \ldots$

(3.4) $\qquad \eta_0^j := 0, \qquad \eta_k^j := \min\{i > \eta_{k-1}^j + [n_j^\alpha] : \mathbf{S}_i \in L\}, \qquad \mathbf{Y}_k^j := \mathbf{S}_{\eta_k^j}.$

Furthermore, let ν_j be the largest integer N for which

(3.5) $\qquad\qquad\qquad\qquad\qquad\qquad \eta_N^j + [n_j^\alpha] \leq n_j.$

Put $f(n_j) = [n_j^{(1-\alpha)(1-\kappa)/2}]$ for any $0 < \kappa < 1$. Since $\mathbf{S}_i \in L$ if and only if $Z_i = 0$, we can apply Fact 6 to get

(3.6) $\qquad\qquad\qquad\qquad\qquad \mathbf{P}\left(\nu_j < f(n_j)\right) \leq C n_j^{(\alpha-1)\kappa/2}.$

Introduce further the events

(3.7) $\qquad\qquad A_k^j = \{\xi^{(2)}(\mathbf{Y}_k^j, \eta_k^j + [n_j^\alpha]) - \xi^{(2)}(\mathbf{Y}_k^j, \eta_k^j) < K(\log n_{j+1})^2\}.$

For fixed j the events A_k^j are independent in k and having the same probability as A_1^j. Using the above notations and (3.6)

(3.8) $\quad \mathbf{P}\left(\xi_L^{(2)}(n_j) < K(\log n_{j+1})^2\right) \leq \mathbf{P}\left(\bigcap_{k=1}^{\nu_j} A_k^j\right) \leq C n_j^{(\alpha-1)\kappa/2} + \left(\mathbf{P}(A_1^j)\right)^{f(n_j)}.$

Now by Fact 8 we have

$$\mathbf{P}(A_1^j) = \mathbf{P}(\xi^{(2)}(\mathbf{0}, [n_j^\alpha]) < K(\log n_{j+1})^2)$$

(3.9) $$\leq 1 - \exp\left\{-\frac{K\pi}{\alpha}(1+\delta)(\beta \log j)\left(\frac{\log(j+1)}{\log j}\right)^2\right\}.$$

Consequently, we have for j big enough

$$\mathbf{P}\left(\xi_L^{(2)}(n_j) < K(\log n_{j+1})^2\right)$$

$$\leq C n_j^{(\alpha-1)\kappa/2} + \left(1 - \exp\left\{-\frac{K\pi}{\alpha}(1+\delta)^2(\beta \log j)\right\}\right)^{(n_j)^{\frac{(1-\alpha)(1-\kappa)}{2}}}$$

$$\leq C j^{\beta(\alpha-1)\kappa/2} + \left(1 - \frac{1}{j^{\frac{K\pi}{\alpha}\beta(1+\delta)^2}}\right)^{j^{\frac{\beta(1-\alpha)(1-\kappa)}{2}}}$$

(3.10) $$\leq C j^{\beta(\alpha-1)\kappa/2} + C \exp\{-j^{\frac{\beta(1-\alpha)(1-\kappa)}{2} - \frac{K\pi}{\alpha}\beta(1+\delta)^2}\}.$$

17

For given $0 < \alpha < 1$ and $\kappa > 0$ select an integer β such that

$$\beta(\alpha - 1)\kappa/2 < -1$$

so that the first term in (3.10) is summable in j. The second term will be summable in j whenever

$$(3.11) \qquad \frac{(1-\alpha)(1-\kappa)}{2} > \frac{K\pi}{\alpha}(1+\delta)^2.$$

On choosing $\alpha = 1/2$, $\delta > 0$ and $\kappa > 0$ small, we conclude that (3.10) is summable in j if

$$(3.12) \qquad K < \frac{1}{8\pi}.$$

Using again Borel-Cantelli lemma and the usual monotonicity argument, we get

$$(3.13) \qquad \limsup_{n\to\infty} \frac{\xi_L^{(2)}(n)}{(\log n)^2} \geq \frac{1}{8\pi} \qquad a.s.,$$

proving our theorem. $\square$

Proof of Theorem 1.2. We will only prove the lower bound (see Remark after the theorem). First observe that the condition $\alpha \leq 1/2$ is not a real restriction as by the LIL (Fact 1) with probability 1 the walk in the time interval $[0, n]$ remains in $B(r_n)$ for $r_n = n^\alpha$, $\alpha > 1/2$, n large, thus $\xi_{B(r_n)}^{(2)}(n) = \xi^{(2)}(n)$ eventually with probability 1.

In case $\alpha \leq 1/2$, using the LIL again, during the first $[r_n^{2-\varepsilon}]$ steps the walk remains in $B(r_n)$ with probability 1 for any $\varepsilon > 0$ and for all but finitely many n. Hence by Theorem C for any $\varepsilon > 0$ with probability 1 eventually,

$$
\begin{aligned}
\xi_{B(r_n)}^{(2)}(n) &\geq \xi^{(2)}([r_n^{2-\varepsilon}]) \geq \frac{1}{\pi}\left(\log r_n^{2-\varepsilon}\right)^2 (1-\varepsilon) \\
(3.14) \qquad &= \frac{1}{\pi}\left(\log n^{\alpha(2-\varepsilon)}\right)^2 (1-\varepsilon) = \frac{\alpha^2(2-\varepsilon)^2}{\pi}(\log n)^2(1-\varepsilon).
\end{aligned}
$$

Let $\varepsilon \to 0$ to get

$$(3.15) \qquad \liminf_{n\to\infty} \frac{\xi_{B(r_n)}^{(2)}(n)}{(\log n)^2} \geq \frac{4\alpha^2}{\pi} \qquad a.s.$$

Thus we have our theorem. $\square$

Remark. Observe that in case $\alpha = 1/2$ the lim inf and lim sup coincide.

Proof of Theorem 1.3. Let $r_n = \exp\{(\log n)^\beta\}$. To get the lower half of the theorem, by the same argument as in the proof of Theorem 1.2, in the first $[r_n^{2-\varepsilon}]$ steps the walk remains in $B(r_n)$ with probability 1, thus by Theorem C we have with probability 1 eventually

$$
\begin{aligned}
\xi_{B(r_n)}^{(2)}(n) &\geq \xi^{(2)}([r_n^{2-\varepsilon}]) \geq \frac{1}{\pi}\left(\log r_n^{2-\varepsilon}\right)^2 (1-\varepsilon) \\
&= \frac{1}{\pi}(\log n)^{2\beta}(2-\varepsilon)^2(1-\varepsilon)
\end{aligned}
$$

(3.16)

Sending now $\varepsilon \to 0$ we get

(3.17)
$$
\liminf_{n\to\infty} \frac{\xi_{B(r_n)}^{(2)}(n)}{(\log n)^{2\beta}} \geq \frac{4}{\pi} \qquad a.s.
$$

To get the other half of the theorem, consider first a disc $A_n := B(\exp\{n^\gamma\})$ where $1/2 \leq \gamma < 1$. Recall the definition of ρ_n in (2.6). First we show that for arbitrary small $\delta > 0$

(3.18)
$$
\max_{\mathbf{x}\in A_n} \xi^{(2)}(\mathbf{x},\rho_n) \leq n^{2\gamma+\delta} \qquad a.s.
$$

for all but finitely many n.

Using the definition of $Y_n(\mathbf{x})$ in Fact 3 and Lemma 2.1 we have that for any $u > 2$

$$
\begin{aligned}
\mathbf{P}(\max_{\mathbf{x}\in A_n} \xi^{(2)}(\mathbf{x},\rho_n) > u\,n) &\leq \mathbf{P}(\max_{\mathbf{x}\in A_n} \sum_{k=1}^n Y_k(\mathbf{x}) > u\,n) \\
&\leq C_1 \exp\{2n^\gamma\} \max_{\mathbf{x}\in A_n} \mathbf{P}(\sum_{k=1}^n Y_k(\mathbf{x}) > u\,n) \\
&\leq C_1 \exp\{2n^\gamma\} \max_{\mathbf{x}\in A_n} \exp\{np(\mathbf{x})(1-u/2)\} \\
&\leq C_1 \exp\{2n^\gamma\} \exp\{C_2 n^{1-\gamma}(1-u/2)\},
\end{aligned}
$$

(3.19)

where in the last inequality we used that for all $\mathbf{x} \in A_n$ $p(\mathbf{x}) \geq C_2 n^{-\gamma}$ by Fact 2. Selecting now $u = n^\theta$ we get from (3.19) that for n big enough

(3.20)
$$
\mathbf{P}(\max_{\mathbf{x}\in A_n} \xi^{(2)}(\mathbf{x},\rho_n) > n^{1+\theta}) \leq C_1 \exp\{2n^\gamma\} \exp\{-C_3 n^{1+\theta-\gamma}\}.
$$

The probabilities in (3.20) can be summed up for n if $1+\theta-\gamma > \gamma$. Thus for any $1/2 \leq \gamma < 1$ and for an arbitrary small $\delta > 0$ we might select $\theta > 0$ such that

19

$$(3.21) \qquad\qquad\qquad\qquad 1 + \theta = 2\gamma + \delta$$

to get by Borel-Cantelli lemma that (3.18) holds.

Applying now (3.18) with $n = \xi^{(2)}(\mathbf{0}, k)$ we get that

$$
\begin{aligned}
\max_{\mathbf{x} \in A_{\xi^{(2)}(\mathbf{0},k)}} \xi^{(2)}(\mathbf{x}, k) \;&\le\; \max_{\mathbf{x} \in A_{\xi^{(2)}(\mathbf{0},k)+1}} \xi^{(2)}\!\left(\mathbf{x}, \rho_{\xi^{(2)}(\mathbf{0},k)+1}\right) \\
(3.22) \qquad\qquad\qquad &\le\; \left(\xi^{(2)}(\mathbf{0}, k) + 1\right)^{2\gamma+\delta} \le (\log k)^{2\gamma+\delta'} \qquad a.s.
\end{aligned}
$$

for any $\delta' > \delta$ and for all but finitely many k by Theorem A. On the other hand, (3.22) and Theorem A imply that for any $0 < \gamma' < \gamma$

$$(3.23) \qquad \max_{\mathbf{x} \in B(\exp\{(\log k)^{\gamma'}\})} \xi^{(2)}(\mathbf{x}, k) \le \max_{\mathbf{x} \in A_{\xi^{(2)}(\mathbf{0},k)}} \xi^{(2)}(\mathbf{x}, k) \le (\log k)^{2\gamma+\delta'} \qquad a.s.$$

for all but finitely many k, which is equivalent to our statement. $\square$

Proof of Corollary 1.1. Clearly the upper bound for $\xi^{(2)}_{B(r_n)}(n)$ or $\xi^{(2)}_L(n)$ holds for $\xi^{(2)}_{B(r_n)} \cap L$ as well. On the other hand, to get the lower bounds, observe that with probability one, in the first $[r_n^{2-\varepsilon}]$ steps the walk remains in $B(r_n)$ eventually. Consequently

$$(3.24) \qquad\qquad \xi^{(2)}_L(r_n^{2-\varepsilon}) = \xi^{(2)}_{L \cap B(r_n)}(r_n^{2-\varepsilon}) \le \xi^{(2)}_{L \cap B(r_n)}(n)$$

for all but finitely many n. But (3.24) and Theorem 1.1 imply our statements. $\square$

4. Proofs of the higher dimensional results.

Now let $\{\mathbf{S}_n\}$ be a d-dimensional ($d \ge 3$) simple symmetric random walk as defined in the Introduction with steps $\{\mathbf{X}_i\}$.

Proof of Theorem 1.4. The proof of this theorem is very similar to the proof of Theorem 1.1. Consider the subspace

$$S_{d-1} = \left\{\mathbf{x} \in \mathcal{Z}_d : a_1 x_1 + a_2 x_2 + ... + a_d x_d = 0\right\}$$

with integer coefficients $a_1, a_2, ...a_d$, not all of them zero. Without loss of generality we may assume that the largest common divisor of $(a_1, \ldots, a_d)$ is equal to 1.

Define a one-dimensional random walk with the following steps:

$$Y_i = a_\ell \quad \text{if} \quad \mathbf{X}_i = \mathbf{e}_\ell, \qquad \ell = 1, 2, ...d, \qquad i = 1, 2, ...$$

(4.1)
$$Y_i = -a_\ell \quad \text{if} \quad \mathbf{X}_i = -\mathbf{e}_\ell, \qquad \ell = 1, 2, ...d, \qquad i = 1, 2, ...$$

Note that these values are not necessarily distinct. In that case we sum up the probabilities $1/(2d)$ according to their multiplicity.

$Z_n = \sum_{i=1}^n Y_i$, $n = 1, 2, \ldots$ is an aperiodic one-dimensional symmetric random walk with $Z_n = 0$ if and only if $\mathbf{S}_n \in S_{d-1}$. Thus denoting by $V^{S_{d-1}}(n) = \#\{i : 1 \le i \le n, \ \mathbf{S}_i \in S_{d-1}\}$, the number of visits of $\{\mathbf{S}_i\}$ up to time n in S_{d-1}, we have

(4.2)
$$V^{S_{d-1}}(n) = \xi^Z(0, n),$$

where $\xi^Z(0, n)$ is the local time at zero up to time n of the random walk $\{Z_i\}$.

To get the upper bound in our theorem, denote by $D(n)$ the set of lattice points on S_{d-1} which are visited by $\{\mathbf{S}_i\}$ up to time n and denote $|D(n)|$ the number of points in $D(n)$. Select a subsequence $n_j = [e^j]$, $j = 1, 2, \ldots$ and define the events for any $\mathbf{x} \in \mathcal{Z}_d$

$$C_\mathbf{x}^j = \{\xi^{(d)}(\mathbf{x}, n_{j+1}) > \lambda(\log n_j)\}.$$

Then using the exact distribution (1.13), and Fact 9 we conclude, that

$$\mathbf{P}(\xi_{S_{d-1}}^{(d)}(n_{j+1}) > \lambda \log n_j)$$

$$\le \mathbf{P}\left(\bigcup_{\mathbf{x} \in D(n_{j+1})} C_\mathbf{x}^j, \ |D(n_{j+1})| \le \sqrt{n_{j+1} \log n_{j+1}}\right) + \mathbf{P}\left(|D(n_{j+1})| > \sqrt{n_{j+1} \log n_{j+1}}\right)$$

$$\le \mathbf{P}\left(\bigcup_{\mathbf{x} \in D(n_{j+1})} C_\mathbf{x}^j, \ |D(n_{j+1})| \le \sqrt{n_{j+1} \log n_{j+1}}\right) + \mathbf{P}\left(\xi^Z(0, n_{j+1}) > \sqrt{n_{j+1} \log n_{j+1}}\right)$$

$$\le \sqrt{n_{j+1} \log n_{j+1}} \ \mathbf{P}(\xi^{(d)}(\mathbf{0}, n_{j+1}) > \lambda \log n_j) + C_2 \exp\left(-(j+1)\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

$$\le \sqrt{n_{j+1} \log n_{j+1}} \ \mathbf{P}(\xi^{(d)}(\mathbf{0}, \infty) > \lambda \log n_j) + C_2 \exp\left(-(j+1)\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

$$\le \sqrt{n_{j+1} \log n_{j+1}} \exp\left(-\frac{\lambda}{\lambda_d} j\right) + C_2 \exp\left(-(j+1)\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

(4.3)
$$\le C\sqrt{j} \exp\left(j\left(\frac{1}{2} - \frac{\lambda}{\lambda_d}\right)\right) + C_2 \exp\left(-(j+1)\frac{(1-\varepsilon)\sigma^2}{2}\right)$$

21

for j big enough, where we used in the second inequality above that $|D(n)| \leq V^{S_{d-1}}(n)$. The σ above depends only on $a_1, a_2, ...a_d$. It is easy to see that the last line of (4.3) is summable in j whenever $\lambda > \lambda_d/2$, which in turn, using Borel-Cantelli lemma and the usual monotonicity argument implies

$$(4.4) \qquad \limsup_{n \to \infty} \frac{\xi^{(2)}_{S_{d-1}}(n)}{\log n} \leq \frac{\lambda_d}{2} \qquad a.s.,$$

so we have the upper half of the theorem.

To get the lower bound, we will follow Erdős and Taylor's argument again with the appropriate modifications. We consider the walk $\{\mathbf{S}_n\}$, wait $[(\log n)^2]$ steps. After that many steps we wait until the walk arrives back to S_{d-1}. Then wait again $[(\log n)^2]$ steps, and repeat this procedure over and over again. The probability that an arrival point to S_{d-1} will be visited by the walk again in the next $[\log n]$ steps is by (1.12)

$$1 - \gamma_d + O\left(\frac{1}{(\log n)^{1/2}}\right)$$

and the probability that it will be visited at least $[\lambda \log n]$ times (with some $\lambda < 1$) within $[(\log n)^2]$ steps is greater than

$$(4.5) \qquad \left(1 - \gamma_d + O\left(\frac{1}{(\log n)^{1/2}}\right)\right)^{[\lambda \log n]}.$$

These ideas will be combined with our Fact 6, and applied for subsequences, just like in the proof of Theorem 1.1.

Let $n_j = j^\beta$, $\quad j = 1, 2, \ldots$ with some integer β. Define

$$(4.6) \qquad \eta_0^j := 0, \qquad \eta_k^j := \min\{i > \eta_{k-1}^j + [(\log n_j)^2] : \mathbf{S}_i \in S_{d-1}\}, \qquad \mathbf{Y}_k^j := \mathbf{S}_{\eta_k^j}.$$

Furthermore, let ν_j be the largest integer N for which

$$(4.7) \qquad \eta_N^j + [(\log n_j)^2] \leq n_j.$$

Selecting for any $0 < \delta < 1$, $f(n_j) = \left[\left(\frac{n_j}{(\log n_j)^2}\right)^{(1-\delta)/2}\right]$, it is easy to see that from Fact 6 we get

$$(4.8) \qquad \mathbf{P}\left(\nu_j < f(n_j)\right) \leq C \left(\frac{(\log n_j)^2}{n_j}\right)^{\delta/2}.$$

22

Introduce further the events

(4.9) $\qquad A_k^j = \{\xi^{(d)}(\mathbf{Y}_k^j, \eta_k^j + [(\log n_j)^2]) - \xi^{(d)}(\mathbf{Y}_k^j, \eta_k^j) < \lambda \log n_{j+1}\}.$

For fixed j the events A_k^j are independent in k and having the same probability as A_1^j. Using (4.8) we get that

$$\mathbf{P}\left(\xi_{S_{d-1}}^{(d)}(n_j) < \lambda \log n_{j+1}\right) \leq \mathbf{P}\left(\bigcap_{k=1}^{\nu_j} A_k^j\right)$$

(4.10) $$\leq C\left(\frac{(\log n_j)^2}{n_j}\right)^{\delta/2} + \left(\mathbf{P}(A_1^j)\right)^{f(n_j)}.$$

Now by (4.5) we have

$$\mathbf{P}(A_1^j) = \mathbf{P}(\xi^{(d)}(\mathbf{0}, [(\log n_j)^2]) < \lambda(\log n_{j+1}))$$

(4.11) $$< 1 - \left(1 - \gamma_d + O\left(\frac{1}{(\log n_j)^{1/2}}\right)\right)^{\lambda \log n_{j+1}} = 1 - n_{j+1}^{-\frac{\lambda}{\lambda_d}\left(1 + O\left(\frac{1}{(\log n_j)^{1/2}}\right)\right)}.$$

Consequently,

$$\mathbf{P}\left(\xi^{(d)}(n_j) < \lambda(\log n_{j+1})\right)$$

$$\leq C\left(\frac{\beta^2(\log j)^2}{j^\beta}\right)^{\delta/2} + \left(1 - n_{j+1}^{-\frac{\lambda}{\lambda_d}\left(1 + O\left(\frac{1}{(\log n_j)^{1/2}}\right)\right)}\right)^{\left[\left(\frac{n_j}{(\log n_j)^2}\right)^{\frac{1-\delta}{2}}\right]}$$

$$\leq j^{-\frac{\beta\delta}{3}} + \left(1 - (j+1)^{-\beta\frac{\lambda}{\lambda_d}\left(1 + O\left(\frac{1}{(\log j)^{1/2}}\right)\right)}\right)^{j^{\frac{\beta(1-2\delta)}{2}}}$$

(4.12) $$\leq j^{-\frac{\beta\delta}{3}} + \exp\left(-\frac{1}{2}j^{\beta\left(\frac{1}{2}-\delta-\frac{\lambda}{\lambda_d}+O\left(\frac{1}{(\log j)^{1/2}}\right)\right)}\right)$$

for $0 < \delta < 1/2$ and j big enough. It is easy to see that in (4.12) selecting $\delta > 0$ as small as necessary the second term is summable in j if $\lambda < \lambda_d/2$. On the other hand, one can select a β big enough (depending on δ) that the first term is summable in j. Borel-Cantelli lemma and the usual monotonicity argument results in

(4.13) $$\liminf_{n \to \infty} \frac{\xi_{S_{d-1}}^{(2)}(n)}{\log n} \geq \frac{\lambda_d}{2} \qquad a.s.$$

23

which combined with (4.4) implies our theorem. □

Proof of Theorem 1.5. Recall that

$$S_{d-2} = \{\mathbf{x} \in \mathbb{Z}_d : a_1 x_1 + a_2 x_2 + \ldots + a_d x_d = 0, \quad b_1 x_1 + b_2 x_2 + \ldots + b_d x_d = 0\}$$

with integers a_i, b_j. Without loss of generality, we may assume that the largest common divisor of both $(a_1, \ldots, a_d)$ and $(b_1, \ldots, b_d)$ is equal to 1 and $(a_1, \ldots, a_d) \neq (b_1, \ldots, b_d)$.

Define a two-dimensional random walk associated with our d-dimensional random walk with steps

$$\widetilde{Y}_i = (a_r, b_r) \quad \text{if} \quad \mathbf{X}_i = \mathbf{e}_r, \quad r = 1, 2, \ldots, d,$$

$$\widetilde{Y}_i = -(a_r, b_r) \quad \text{if} \quad \mathbf{X}_i = -\mathbf{e}_r, \quad r = 1, 2, \ldots, d.$$

Note that, as in the previous theorem, these values are not necessarily distinct. In that case we sum up the probabilities $1/(2d)$ according to their multiplicity.

Then

$$\widetilde{Z}_n = \widetilde{Y}_1 + \ldots + \widetilde{Y}_n, \quad n = 1, 2, \ldots$$

is a two-dimensional symmetric recurrent (possibly periodic) random walk in $\mathbb{Z}_2$ with finite variance with the property $\mathbf{S}_n \in S_{d-2}$ if and only if $\widetilde{Z}_n = 0$. Let

$$V^{S_{d-2}}(n) = \#\{i : 1 \leq i \leq n, \mathbf{S}_i \in S_{d-2}\}$$

and

$$\xi^{\widetilde{Z}}(\mathbf{0}, n) = \#\{i : 1 \leq i \leq n, \widetilde{Z}_i = \mathbf{0}\},$$

the local time of $\{\widetilde{Z}\}$ at $\mathbf{0}$ up to time n.

Then we have $V^{S_{d-2}}(n) = \xi^{\widetilde{Z}}(\mathbf{0}, n)$.

Now we define an aperiodic random walk with the same property. If $\{\widetilde{Z}_n\}$ is aperiodic, then let $\widehat{Z}_n = \widetilde{Z}_n$, $n = 0, 1, \ldots$ In the case when $\{\widetilde{Z}_n\}$ happens to be periodic, then following Spitzer [9], pp. 65-66 construct an aperiodic walk in $\mathbb{Z}_2$ as follows. The steps (a_r, b_r), $r = 1, \ldots, d$ are supported on a proper subgroup $G \subset \mathbb{Z}_2$. Since under our assumptions $a_r \neq b_r$ at least for one r, the subgroup is two-dimensional, i.e. there exists a basis of G consisting of two vectors, say $\mathbf{u}$, $\mathbf{v}$. Let (α_r, β_r) be the coordinates of (a_r, b_r) related to this basis, i.e. $(a_r, b_r) = \alpha_r \mathbf{u} + \beta_r \mathbf{v}$. Now define a new two-dimensional random walk with steps

$$\widehat{Y}_i = (\alpha_r, \beta_r) \quad \text{if} \quad \widetilde{Y}_i = (a_r, b_r), \quad r = 1, 2, \ldots, d,$$

$$\widehat{Y}_i = -(\alpha_r, \beta_r) \quad \text{if} \quad \widetilde{Y}_i = -(a_r, b_r), \quad r = 1, 2, \ldots, d.$$

24

Then (cf. [9], p. 65)
$$\widehat{Z}_n = \widehat{Y}_1 + \ldots + \widehat{Y}_n, \quad n = 1, 2, \ldots$$

is a symmetric recurrent aperiodic random walk in $\mathcal{Z}_2$. Obviously $\widetilde{Z}_n = \mathbf{0}$ if and only if $\widehat{Z}_n = \mathbf{0}$. Hence we have also $V^{S_{d-2}}(n) = \xi^{\widehat{Z}}(\mathbf{0}, n)$ with obvious notation for the local time of $\{\widehat{Z}\}$.

Now we prove the upper bound in Theorem 1.5, i.e.

(4.14)
$$\limsup_{n \to \infty} \frac{\xi^{(d)}_{S_{d-2}}(n)}{\log \log n} \leq \lambda_d \qquad a.s.$$

In the proof we follow the same lines as in the proof of Theorem 1.4. Choose $\lambda > \lambda_d$ and $\varepsilon > 0$ such that $\delta = \lambda/\lambda_d - 1 - \varepsilon > 0$. Then using (1.13), Fact 5 and Fact 10 we can see as in (4.3) that for n big enough,

$$\mathbf{P}\left(\xi^{(d)}_{S_{d-2}}(n) > u\right) \leq (\log n)^{1+\varepsilon} P(\xi^{(d)}(\mathbf{0}, n) \geq u) + \mathbf{P}(V^{(d-2)}(n) > (\log n)^{1+\varepsilon})$$

$$\leq (\log n)^{1+\varepsilon} \mathbf{P}(\xi^{(d)}(\mathbf{0}, \infty) > u) + \mathbf{P}(\xi^{\widehat{Z}}(\mathbf{0}, n) > (\log n)^{1+\varepsilon})$$

$$\leq (\log n)^{1+\varepsilon} \exp\left(-\frac{u}{\lambda_d}\right) + \exp(-(\log n)^{\varepsilon/2}).$$

Hence choosing $n_j = \left[\exp\left(j^{2/\delta}\right)\right]$,

$$\mathbf{P}(\xi^{(d)}_{S_{d-2}}(n_{j+1}) \geq \lambda \log \log n_j) \leq C(\log n_j)^{-\delta} = \frac{C}{j^2}.$$

Borel-Cantelli lemma and the usual monotonicity argument yields (4.14), since ε can be arbitrary small.

Now we prove the lower bound

(4.15)
$$\liminf_{n \to \infty} \frac{\xi^{(d)}_{S_{d-2}}}{\log \log n} \geq \lambda_d \qquad a.s.$$

Here again we follow the proof of Theorem 1.4. Let $\lambda < \lambda_d$ and $n_j = [\exp(j^{2/(1-\beta)})]$ with some $\lambda/\lambda_d < \beta < 1$ and define

(4.16) $\quad \eta_0^j := 0, \qquad \eta_k^j := \min\{i > \eta_{k-1}^j + [(\log n_j)^2] : \mathbf{S}_i \in S_{d-2}\}, \qquad \mathbf{Y}_k^j := \mathbf{S}_{\eta_k^j}.$

25

Furthermore, let ν_j be the largest integer N for which

(4.17) $$\eta_N^j + [(\log n_j)^2] \leq n_j.$$

Since $\{\widehat{Z}\}$ is a two-dimensional random walk for which Lemma 2.4 holds, choosing $b = [(\log n_j)^2]$, $f(n_j) = (\log n_j)^\beta$ there, we get

$$\mathbf{P}(\nu_j \leq (\log n_j)^\beta) \leq \frac{c \log \log n_j}{(\log n_j)^{1-\beta}}.$$

Similarly to (4.5), the probability that a point will be visited at least $\lambda \log \log n_{j+1}$ times within $[(\log n_j)^2]$ steps is greater than

$$\left(1 - \gamma_d + O\left(\frac{1}{(\log n_j)^{1/2}}\right)\right)^{[\lambda \log \log n_{j+1}]} = (\log n_{j+1})^{-\lambda/\lambda_d}(1 + o(1)).$$

Define the events

$$A_k^j = \{\xi^{(d)}(\mathbf{S}_{\eta_k^j}, \eta_k^j + [(\log n_j)^2]) - \xi^{(d)}(\mathbf{S}_{\eta_k^j}, \eta_k^j) < \lambda \log \log n_{j+1}\},$$

then clearly A_k^j are independent in k for fixed j, and have the same probability, hence we have for j big enough

$$\mathbf{P}(A_k^j) \leq 1 - \frac{1 + o(1)}{(\log n_{j+1})^{\lambda/\lambda_d}}.$$

Moreover,

$$\mathbf{P}(\xi_{S_{d-2}}^{(d)}(n_j) < \lambda \log \log n_{j+1}) \leq \left(1 - \frac{1 + o(1)}{(\log n_{j+1})^{\lambda/\lambda_d}}\right)^{(\log n_j)^\beta} + \frac{c \log \log n_j}{(\log n_j)^{1-\beta}}$$

$$\leq \exp\left(-(1/2)(\log n_j)^{\beta - \lambda/\lambda_d}\right) + \frac{c \log \log n_j}{(\log n_j)^{1-\beta}}.$$

This is summable in j, hence Borel–Cantelli lemma and monotonicity imply (4.15). This together with (4.14) completes the proof of Theorem 1.5. $\square$

Proof of Theorem 1.6. First we prove the lower bound, i.e.

(4.18) $$2\lambda_d \leq \liminf_{n \to \infty} \frac{\xi_{B(r_n)}^{(d)}(n)}{\log r_n} \qquad a.s.$$

26

For any $\delta > 0$ and large enough n we have $r_n^{2-\delta} \leq n$ and by the law of the iterated logarithm (2.1) we conclude that for any $\delta > 0$, $\{\mathbf{S}_i, 1 \leq i \leq r_n^{2-\delta} \leq n\}$ are all in $B(r_n)$ for large n with probability one, hence by Theorem F we have for any η,

$$\xi_{B(r_n)}^{(d)}(n) \geq \xi^{(d)}(r_n^{2-\delta}) \geq (1-\eta)\lambda_d \log r_n^{2-\delta} = (1-\eta)(2-\delta)\lambda_d \log r_n \qquad a.s.$$

for all but finitely many n. Since η and δ are arbitrary, (4.18) follows.

To show the upper bound

$$(4.19) \qquad \limsup_{n\to\infty} \frac{\xi_{B(r_n)}^{(d)}(n)}{\log r_n} \leq 2\lambda_d \qquad a.s.,$$

note that by (2.2) and (2.3) for any $\varepsilon > 0$ and large n the random walk does not hit points in the ball $B(r_n)$ after $r_n^{2+\varepsilon}$ steps with probability one. Hence by Theorem F

$$\xi_{B(r_n)}^{(d)}(n) \leq \xi^{(d)}(r_n^{2+\varepsilon}) \leq (2+2\varepsilon)\lambda_d \log r_n \qquad a.s.$$

for all but finitely many n. Since ε is arbitrary, (4.19) follows.

This completes the proof of Theorem 1.6. $\square$

5. Further consequences and questions.

5.1 A special one dimensional walk.

Consider a simple symmetric walk in $\mathcal{Z}_2$

$$\mathbf{S}_n^{(2)} = \sum_{k=1}^n \mathbf{X}_k = (S_{n,1}, S_{n,2}) = \left(\sum_{k=1}^n X_{k,1}, \sum_{k=1}^n X_{k,2}\right).$$

Clearly the components $X_{k,1}$ and $X_{k,2}$ are dependent. However it is easy to check that the pair

$$Y_{k,1} = X_{k,1} + X_{k,2}, \qquad Y_{k,2} = X_{k,1} - X_{k,2}$$

are independent, with common distribution

$$\mathbf{P}(Y_{k,1} = \pm 1) = \mathbf{P}(Y_{k,2} = \pm 1) = \frac{1}{2}.$$

Consequently,

27

$$V_n = \sum_{k=1}^{n} Y_{k,1} = S_{n,1} + S_{n,2} \quad \text{and} \quad Z_n = \sum_{k=1}^{n} Y_{k,2} = S_{n,1} - S_{n,2}, \quad n = 1, 2 \ldots$$

are independent simple symmetric random walks.

Consider now the consecutive return times of $\{Z_n\}$ to zero, that is to say let

$$\rho_0 := 0, \quad \rho_k := \min\{i > \rho_{k-1} : Z_i = 0\}, \qquad k = 1, 2, \ldots$$

Now clearly our two-dimensional walk $\mathbf{S}_n^{(2)}$ is on the line $x_1 - x_2 = 0$ if and only if $Z_n = 0$, that is to say at the steps ρ_i, $i = 1, 2, \ldots$ Introducing the i.i.d. sequence $U_i = V_{\rho_i} - V_{\rho_{i-1}}$ $i = 1, 2\ldots$ we get a one-dimensional random walk

$$R_n = V_{\rho_n} = \sum_{i=1}^{n} U_i, \qquad n = 1, 2, \ldots$$

For this walk we have from Spitzer [9], p. 89 that

$$\mathbf{P}(U_1 = 0) = 1 - \frac{2}{\pi},$$

$$\mathbf{P}(U_1 = 2k) = \frac{2}{\pi} \frac{1}{4k^2 - 1}, \qquad k = \pm 1, \pm 2, \ldots$$

R_n is in the domain of attraction of the Cauchy distribution, so we will refer to it as Cauchy walk. Some properties of the Cauchy walk was investigated in Taylor [10]. Here we want to point out that our results have some implications for the local time of the Cauchy walk. As

$$R_n = 2k \Leftrightarrow \mathbf{S}_{\rho_n}^{(2)} = (k, k),$$

we conclude that for the local time of R_n

(5.1) $$\eta(2\ell, n) := \#\{i : 1 \le i \le n, R_i = 2\ell\} = \xi^{(2)}((\ell, \ell), \rho_n).$$

Thus from (5.1) we get that

$$\eta(n) = \max_y \eta(y, n) = \xi_L^{(2)}(\rho_n)$$

where $L = L(1, -1)$ is the line $x_1 - x_2 = 0$.

Taking into account the well-known fact that $\log \rho_n \sim 2 \log n$ (see e.g. [8], p. 115) a simple application of our Theorem 1.1 implies that for the maximal local time of the Cauchy walk we have

$$\frac{1}{2\pi} \le \liminf_{n \to \infty} \frac{\eta(n)}{(\log n)^2} \le \limsup_{n \to \infty} \frac{\eta(n)}{(\log n)^2} \le \frac{2}{\pi} \qquad a.s.$$

5.2 Open questions.

1./ Our methods are not powerful enough to get exact constants in Theorems 1.1, 1.2. We don't have any conjecture whether in these theorems the lim inf and lim sup can be replaced by limit and if so what would be the actual value of those limits. In Theorem 1.3 even the exact order escapes us.

2./ In Theorems 1.1, 1.4 and 1.5 we have lines and subspaces going through the origin. These results remain valid for lines and subspaces having a fixed distance from the origin. However it would be interesting to investigate the maximal local time on lines and subspaces having a distance from the origin $d(n) \to \infty$.

3./ In our theorems we have balls centered at the origin. One might be interested in the maximal local time in balls having a center with distance $d(n) \to \infty$ from the origin.

4./ We discussed subspaces and balls in the theorems but other sets would be just as interesting to be investigated. One possibility would be to investigate angular domains, cones and wedges. E.g. in Theorem 1.4 if we take the wedge in between two planes, then we have to have a transition from $\frac{\lambda_d}{2}$ to λ_d as the angle of the wedge increases from $0 \to 2\pi$. Similarly we might ask that if in Theorem 1.5 we replace the line with a cone (centered at the origin) what kind of transition happens as the cone gets wider, that is to say how does the order change from $\log \log n$ to $\log n$ as the angle of the cone changes from $0 \to \pi$ (from the line to $\mathcal{Z}_3$.) Similar questions can be asked in two dimension as well but those would be more interesting if the exact constant were known in Theorem 1.1.

5./ One might ponder that how important is the actual shape of the subset on which the maximum is taken. E.g. it would be interesting to have results on the maximal local time on sets which are not specified in shape just their size are given and of course one has to ensure that they are located close enough to the origin that visits should occur.

6./ All our theorems are about simple symmetric walk. We might ask how important is this restriction. Theorem C is valid for a much wider class of random walks. So we might ask whether our theorems remain valid for aperiodic random walks under certain moment conditions. In Theorem C all the moments has to exist, maybe somewhat less precise results can be ensured even if only the second moment exists. Is it possible to say anything without second moment?

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